

Introduction to Computer Vision

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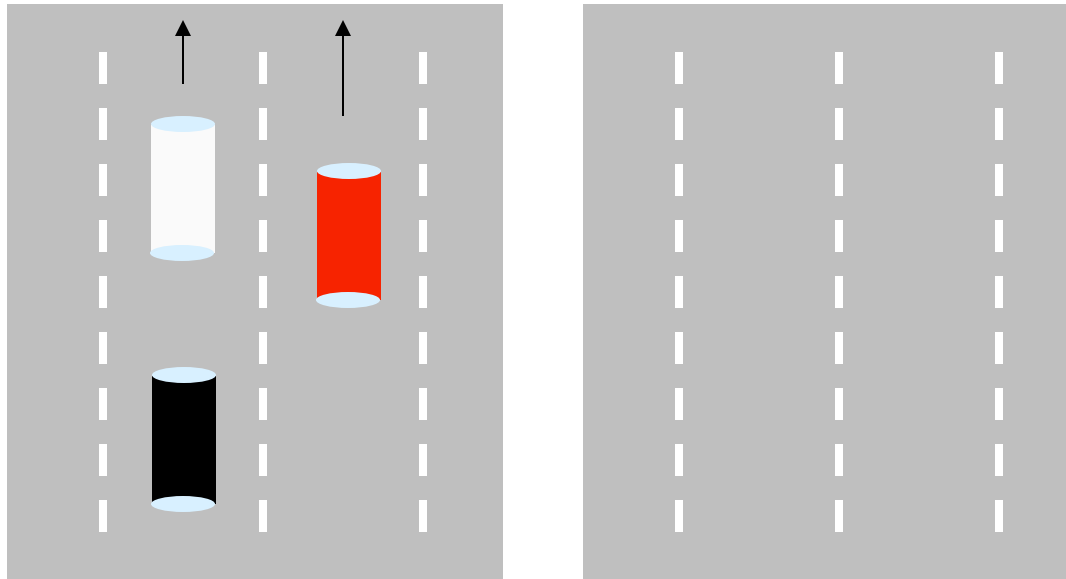
Nov 2009

Tracking and Particle Filtering

Goals

- Today
 - Particle filtering
- Wednesday
 - Binocular stereo
- Friday and beyond
 - Advanced topics – state of the art
 - Short intro to object recognition

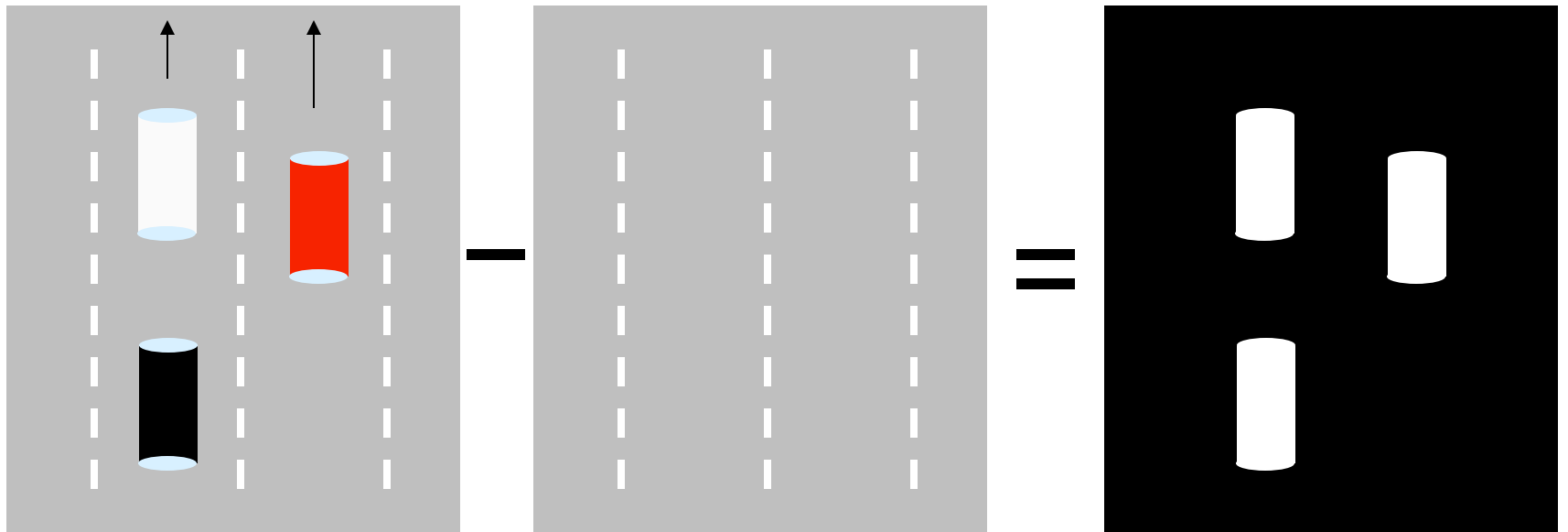
Mathematical Formulation



Goal: estimate car positions at each time instant

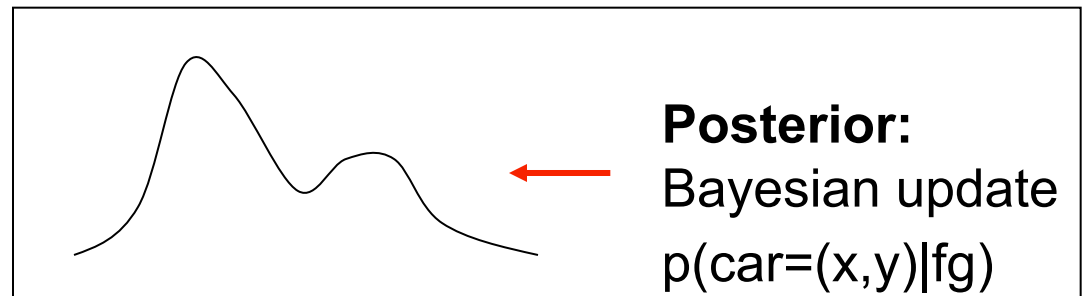
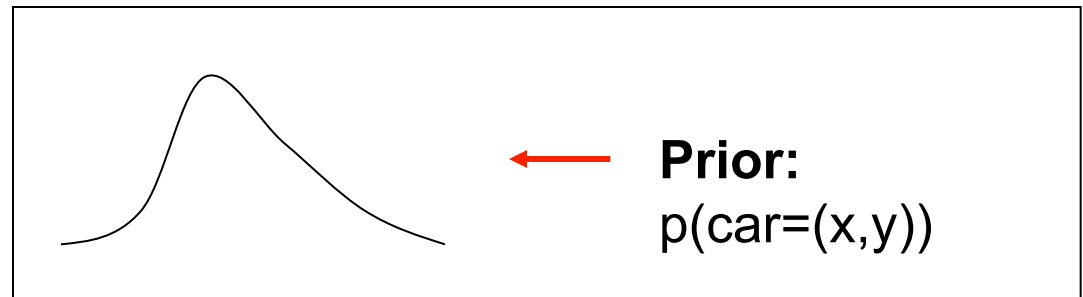
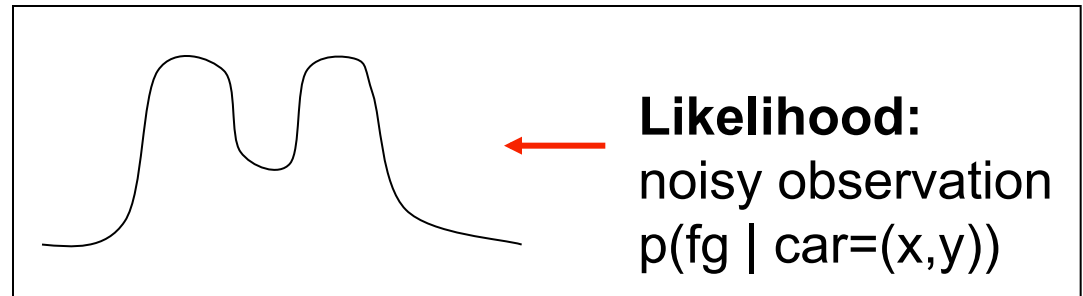
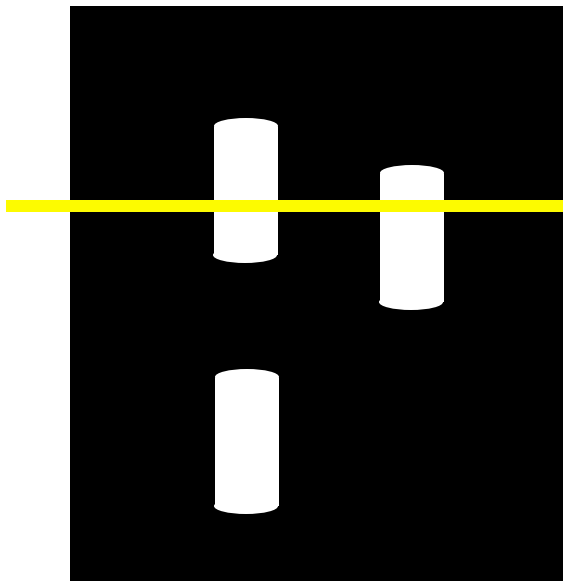
Observations: image sequences and known background

Mathematical Formulation



Define image likelihood: $p(\text{fg} \mid \text{car}=(x,y))$

Mathematical Formulation



system states: car positions

observations: images

Notation

- $\mathbf{x}_k \in \mathbf{R}^d$: internal state at k^{th} frame (hidden random variable, e.g. position of the object in the image).

$\mathbf{X}_k = [\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_k]^T$: history up to time step k

- $\mathbf{z}_k \in \mathbf{R}^c$: measurement at k^{th} frame (observable random variable, e.g. the given image).

$\mathbf{Z}_k = [\mathbf{z}_1, \mathbf{z}_2, \dots, \mathbf{z}_k]^T$:
history up to time step k

Goal

Estimating the posterior probability $p(\mathbf{x}_k | \mathbf{Z}_k)$

How ???

One idea: recursion $p(\mathbf{x}_{k-1} | \mathbf{Z}_{k-1}) \Rightarrow p(\mathbf{x}_k | \mathbf{Z}_k)$

- How to realize the recursion ?
- What assumptions are necessary ?

Recursive Formula: Approximation

$$p(\mathbf{x}_k | \mathbf{Z}_k)$$

$$= p(\mathbf{x}_k | \mathbf{Z}_{k-1}, \mathbf{z}_k)$$

$$\propto p(\mathbf{z}_k | \mathbf{x}_k, \mathbf{Z}_{k-1}) p(\mathbf{x}_k | \mathbf{Z}_{k-1})$$

$$\propto p(\mathbf{z}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{Z}_{k-1})$$

$$\propto p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{Z}_{k-1}) p(\mathbf{x}_{k-1} | \mathbf{Z}_{k-1}) d\mathbf{x}_{k-1}$$

$$\propto p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) p(\mathbf{x}_{k-1} | \mathbf{Z}_{k-1}) d\mathbf{x}_{k-1}$$

Bayes rule:

$$p(a | b) = p(b | a) p(a) / p(b)$$

Assumption:

$$p(\mathbf{z}_k | \mathbf{x}_k, \mathbf{Z}_{k-1}) = p(\mathbf{z}_k | \mathbf{x}_k)$$

Assumption:

$$p(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{Z}_{k-1}) = p(\mathbf{x}_k | \mathbf{x}_{k-1})$$

Integration:

$$p(a) = \int p(a | b) p(b) db$$

Bayesian Formulation

$$p(\mathbf{x}_k | \mathbf{Z}_k) = \kappa p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) p(\mathbf{x}_{k-1} | \mathbf{Z}_{k-1}) d\mathbf{x}_{k-1}$$

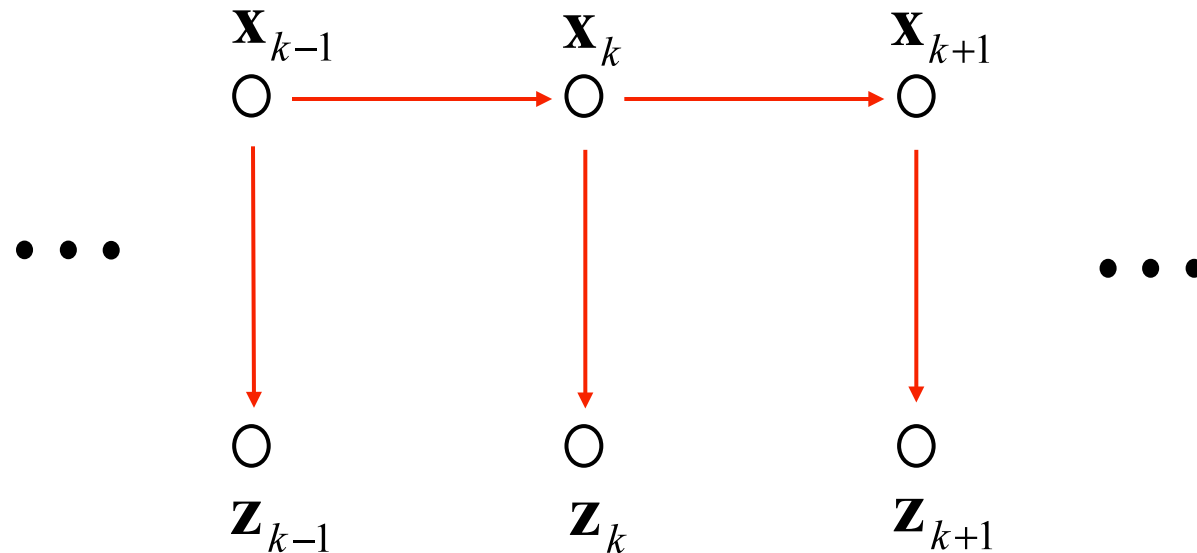
$p(\mathbf{z}_k | \mathbf{x}_k)$: likelihood

$p(\mathbf{x}_k | \mathbf{x}_{k-1})$: temporal prior

$p(\mathbf{x}_{k-1} | \mathbf{Z}_{k-1})$: posterior probability at previous time step

κ : normalizing term

Bayesian Graphical Model



Assumptions:

$$p(\mathbf{z}_k | \mathbf{x}_k, \mathbf{Z}_{k-1}) = p(\mathbf{z}_k | \mathbf{x}_k), \quad p(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{Z}_{k-1}) = p(\mathbf{x}_k | \mathbf{x}_{k-1})$$

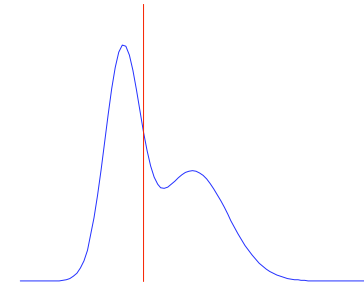
$$p(\mathbf{x}_k | \mathbf{X}_{k-1}) = p(\mathbf{x}_k | \mathbf{x}_{k-1})$$

Estimators

Assume the posterior probability $p(\mathbf{x}_k | \mathbf{Z}_k)$ is known:

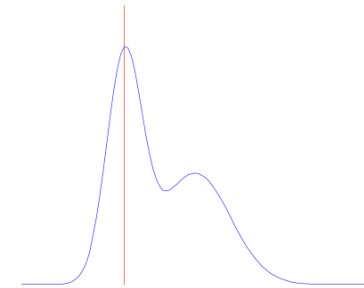
- posterior mean

$$\hat{\mathbf{x}}_k = E(\mathbf{x}_k | \mathbf{Z}_k)$$



- maximum *a posteriori* (MAP)

$$\hat{\mathbf{x}}_k = \arg \max p(\mathbf{x}_k | \mathbf{Z}_k)$$



$p(\mathbf{x}_k | \mathbf{Z}_k)$

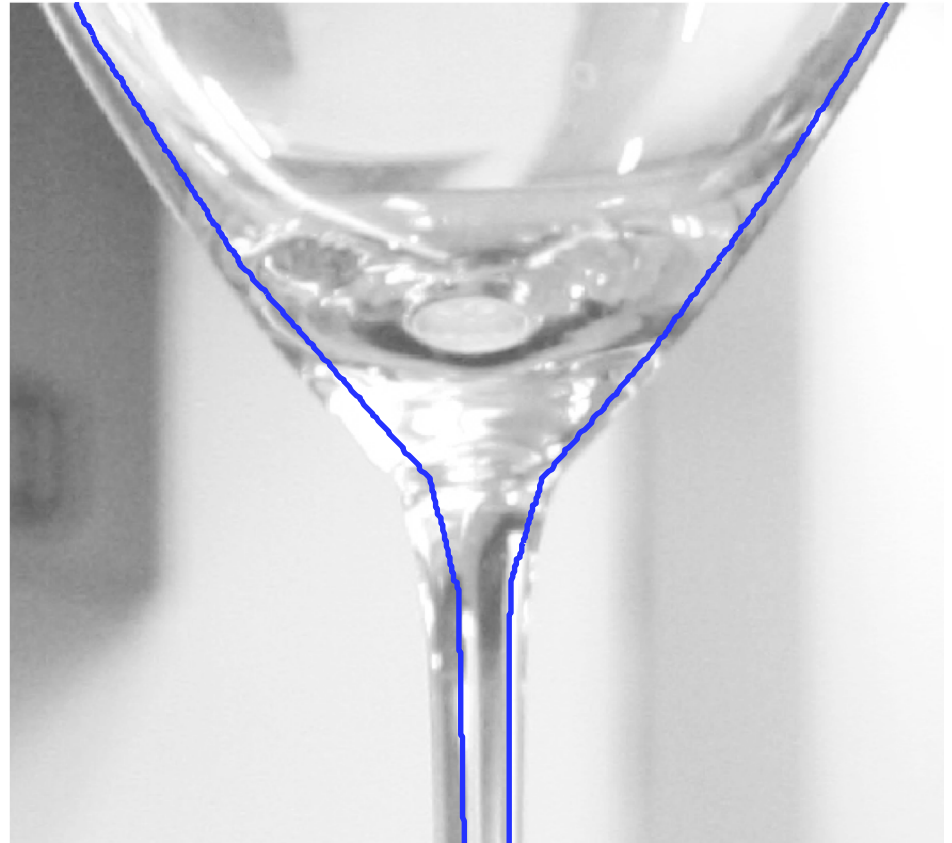
General Model

- $p(\mathbf{x}_k | \mathbf{Z}_k)$ can be an arbitrary, non-Gaussian, multi-modal distribution.
- The recursive equation has no explicit solution, but can be numerically approximated using Monte Carlo techniques.
- If both *likelihood* and *prior* are Gaussian, the solution has closed form and the two estimators (posterior mean & MAP) are the same. Such model is known as the Kalman filter. (Kalman, 1960)

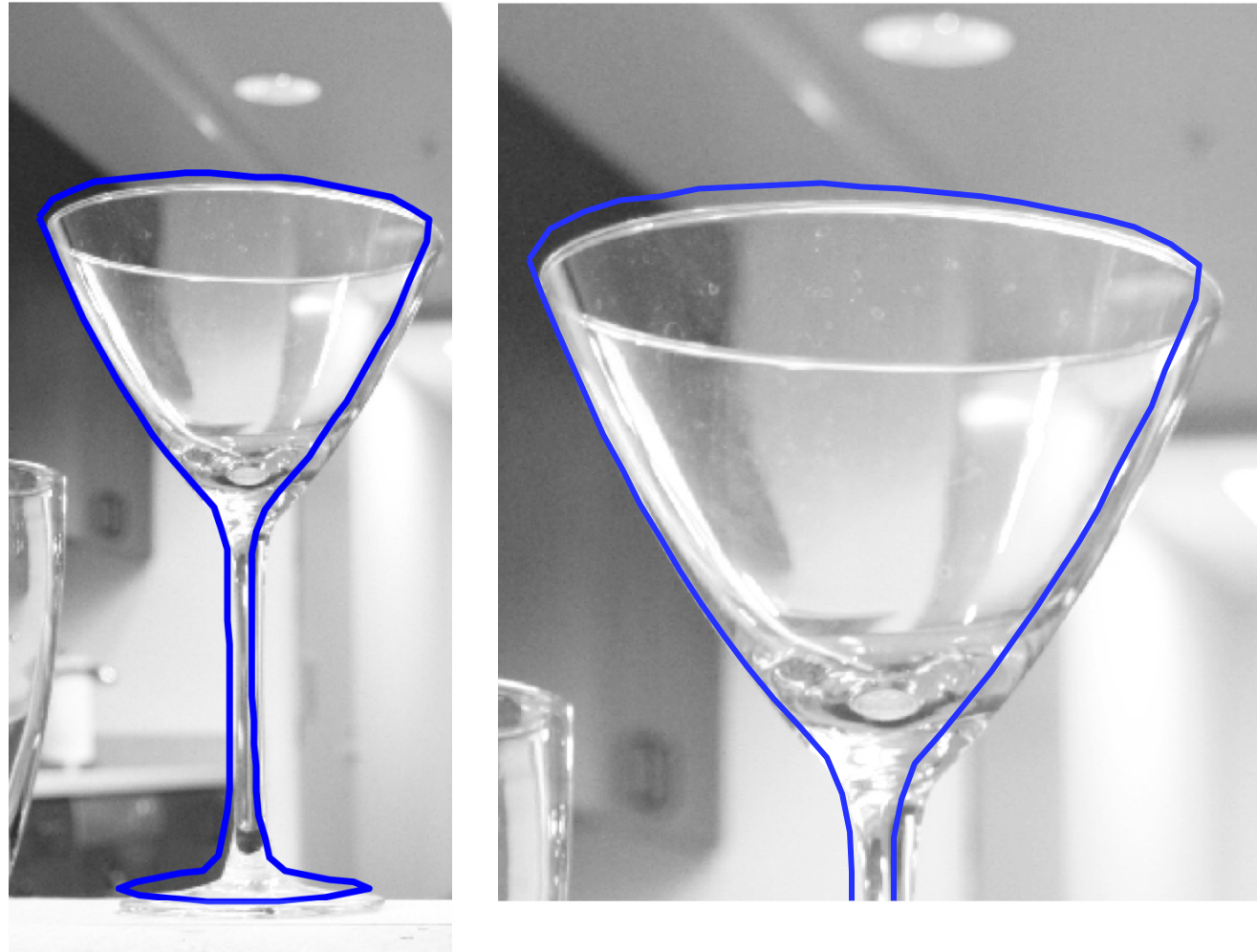
Where's the edge of the glass?



Multi-modal likelihood

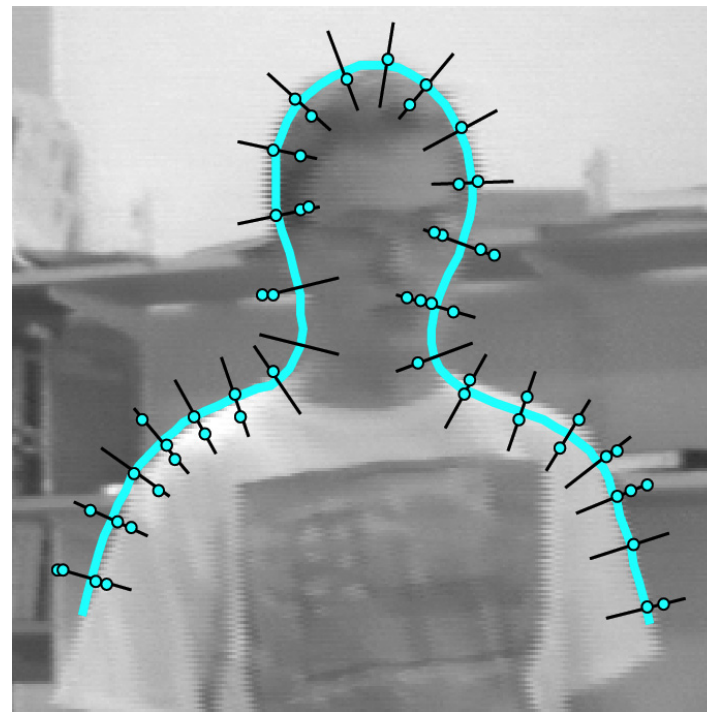


Multi-modal likelihood

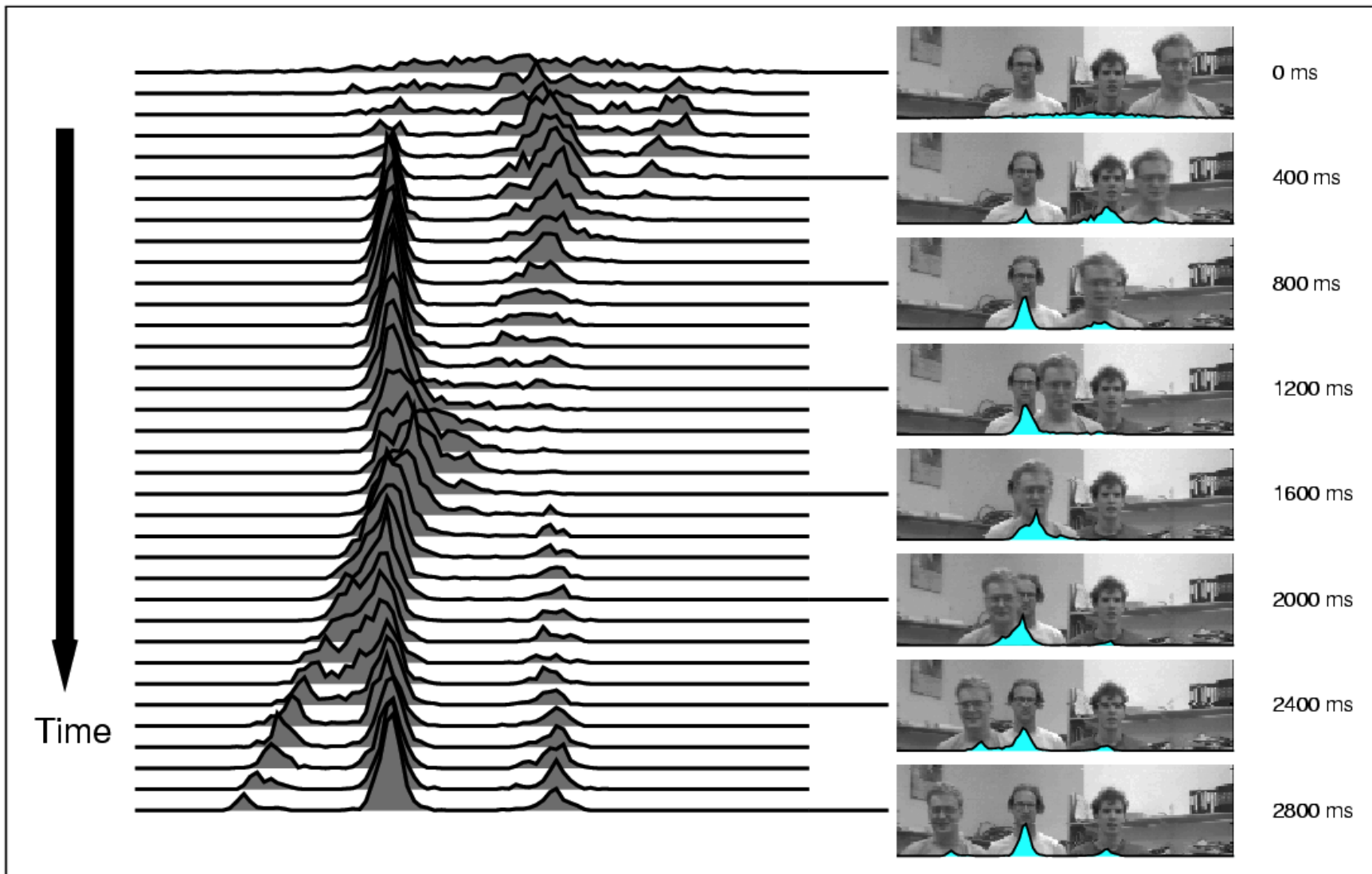


Multi-Modal Likelihoods

Measurement clutter in natural images causes likelihood functions to have multiple, local maxima.

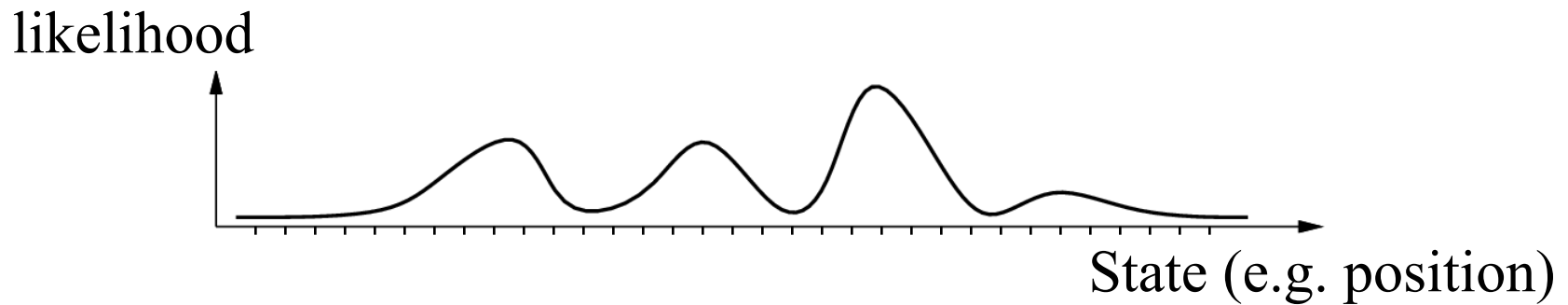


[Isard & Blake, "Condensation - conditional density propagation for visual tracking." IJCV, 1998]



Michael Isard

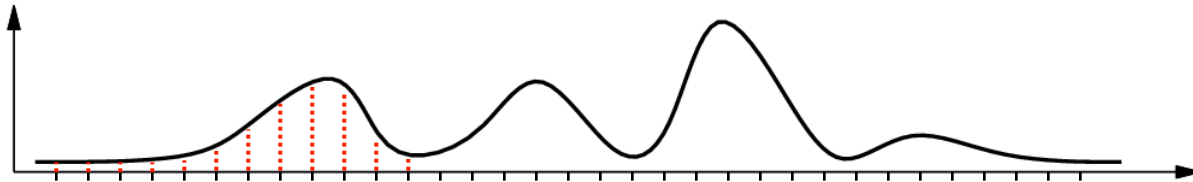
Multi-Modal Likelihood



How can we represent this?

Non-Parametric Approximation

- We could sample at regular intervals



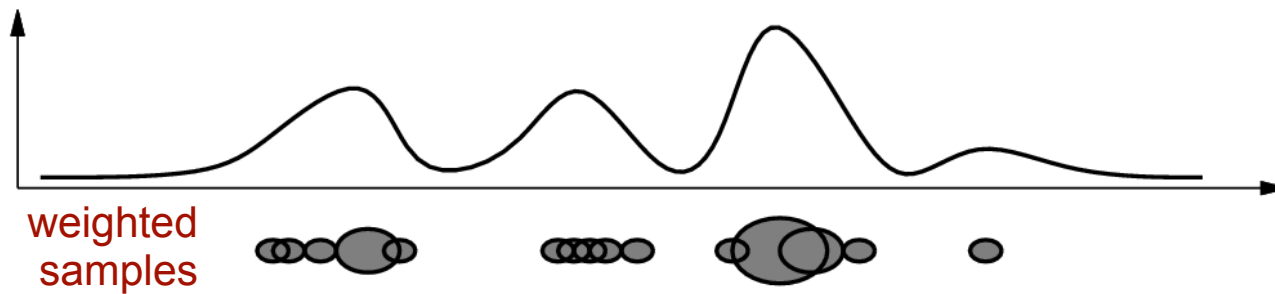
Problems?

most samples have low probability – wasted computation

How finely to discretize

High dimensional space – discretization impractical

Factored Sampling



Weighted samples $S = \{(\mathbf{x}^{(i)}, w^{(i)}); i = 1 \dots N\}$

Normalized likelihood:

$$w_t^{(n)} = \frac{p(\mathbf{z}_t | \mathbf{x}_t^{(n)})}{\sum_{i=1}^N p(\mathbf{z}_t | \mathbf{x}_t^{(i)})}$$

Bayesian Tracking

Posterior over model parameters given an image sequence.

$$p(\mathbf{x}_t | \mathbf{Z}_t) = \int \kappa p(\mathbf{z}_t | \mathbf{x}_t) \left(p(\mathbf{x}_t | \mathbf{x}_{t-1}) p(\mathbf{x}_{t-1} | \mathbf{Z}_{t-1}) \right) d\mathbf{x}_{t-1}$$

Temporal model (prior)

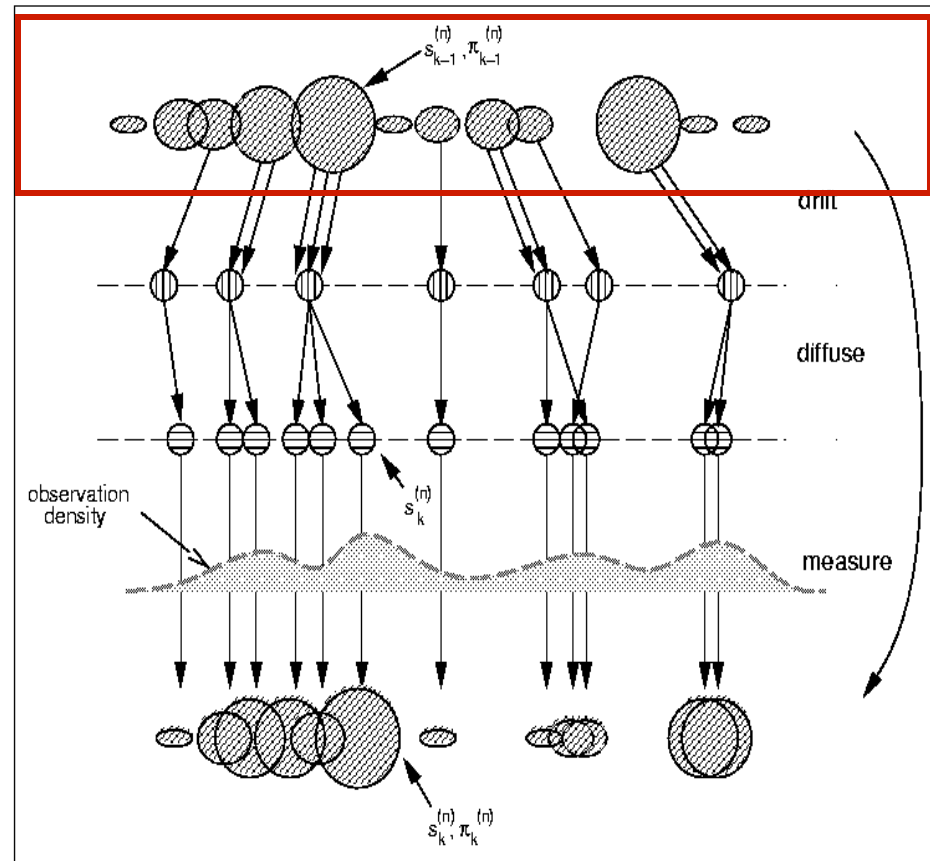
Likelihood of observing the image features given the model parameters

Posterior from previous time instant

Approximate by Monte Carlo sampling

Particle Filter

Posterior $p(\mathbf{x}_{t-1} | \mathbf{Z}_{t-1})$

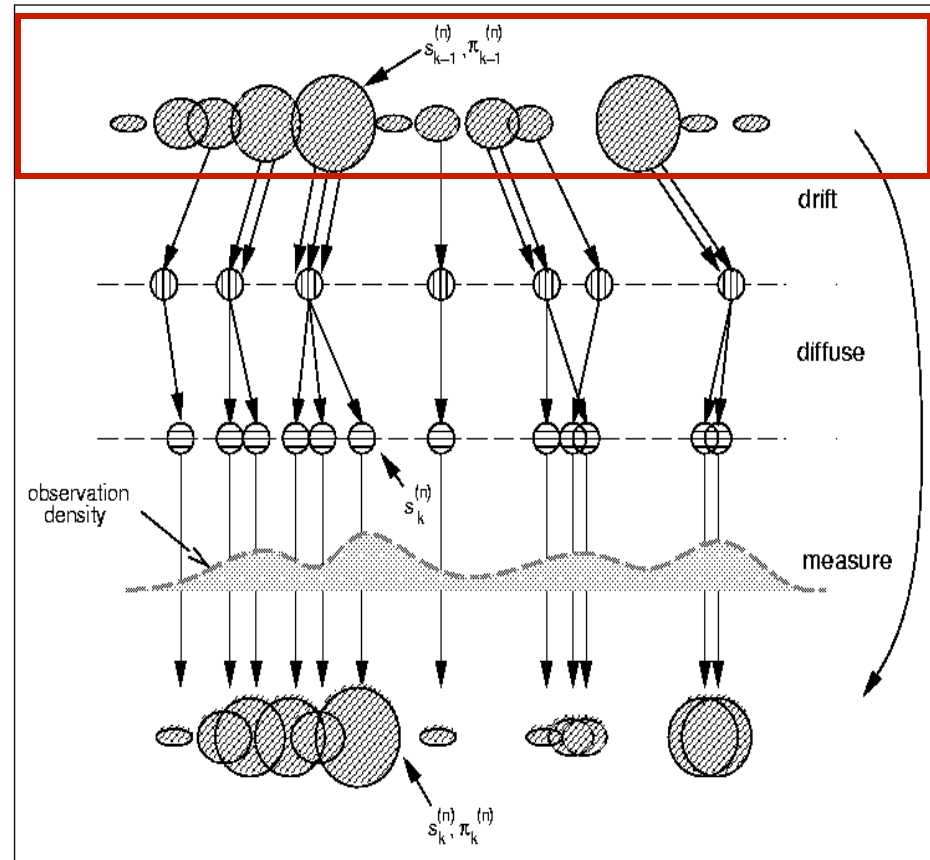


Isard & Blake '96

Particle Filter

Posterior $p(\mathbf{x}_{t-1} | \mathbf{Z}_{t-1})$

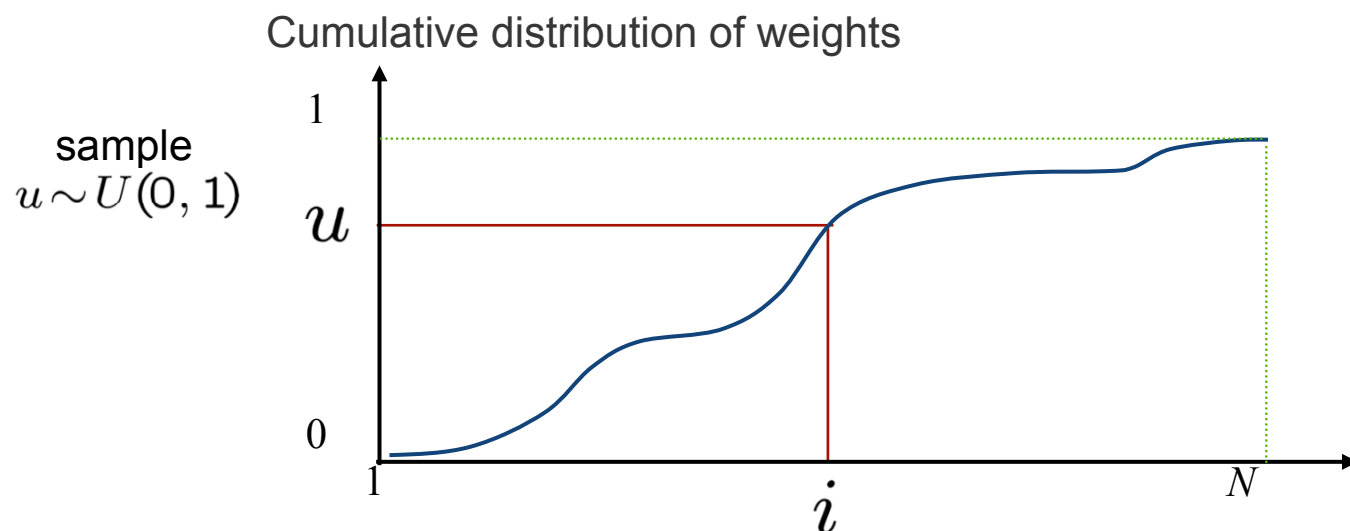
sample



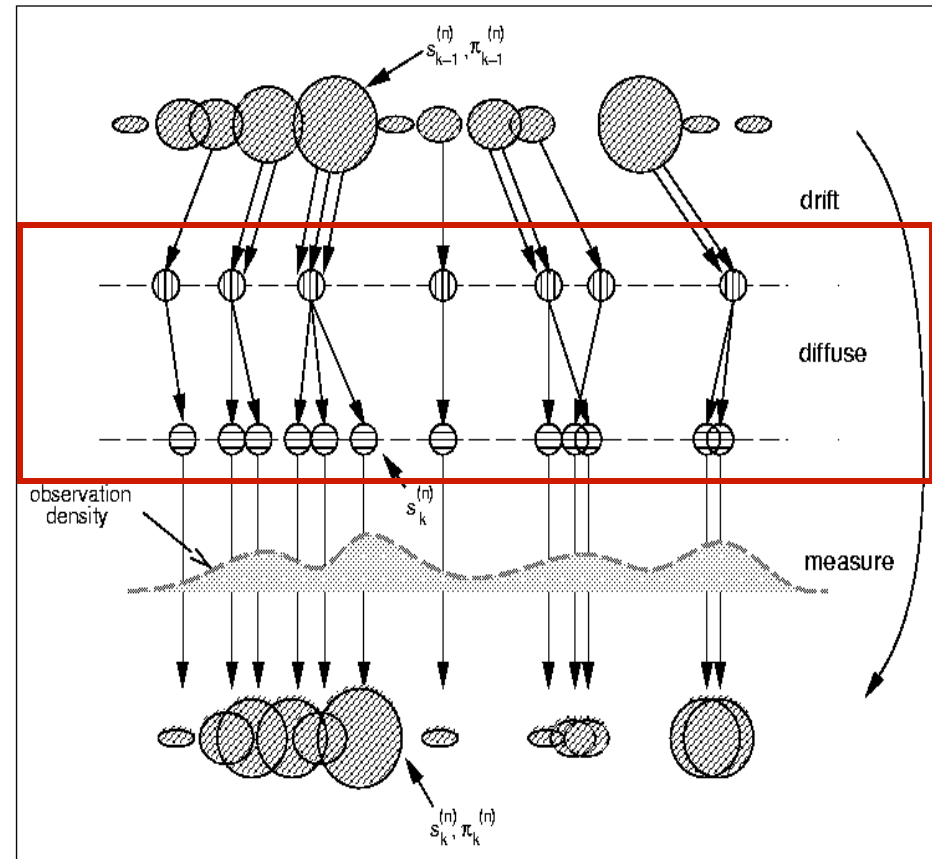
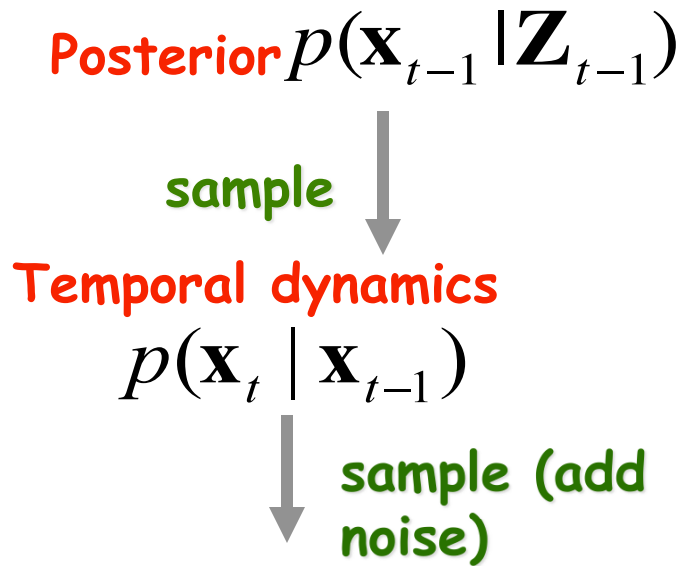
Isard & Blake '96

Monte-Carlo Sampling

Given a weighted sample set $S = \{(\mathbf{x}^{(i)}, w^{(i)}); i = 1 \dots N\}$

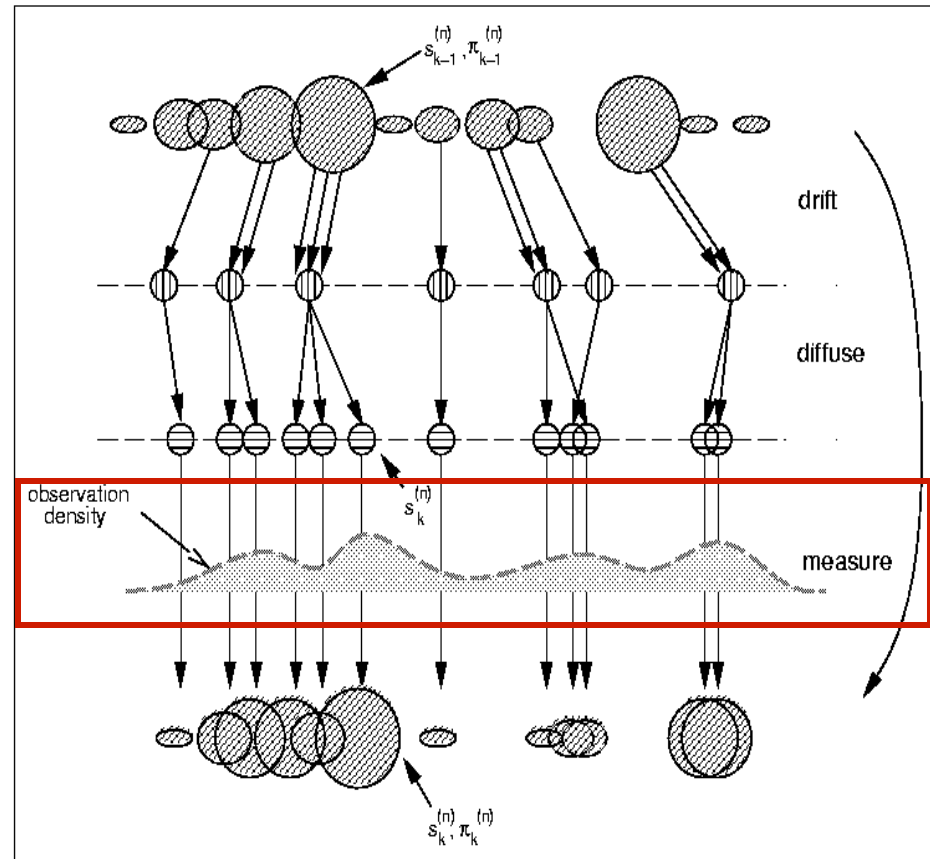
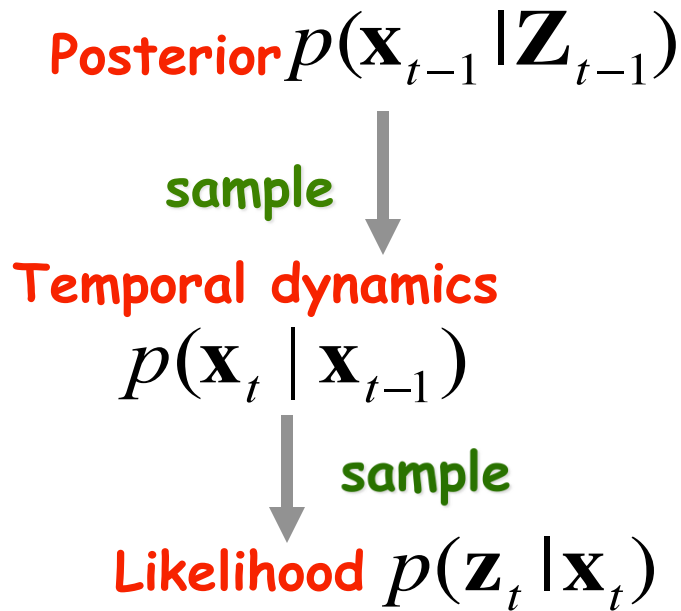


Particle Filter



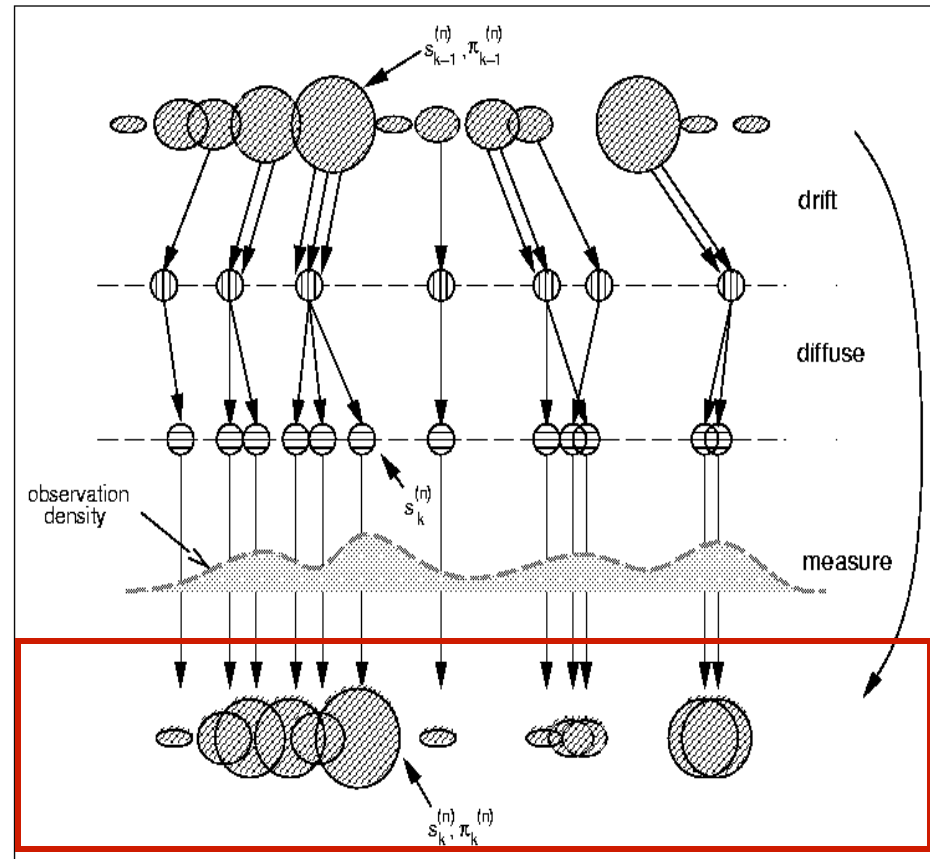
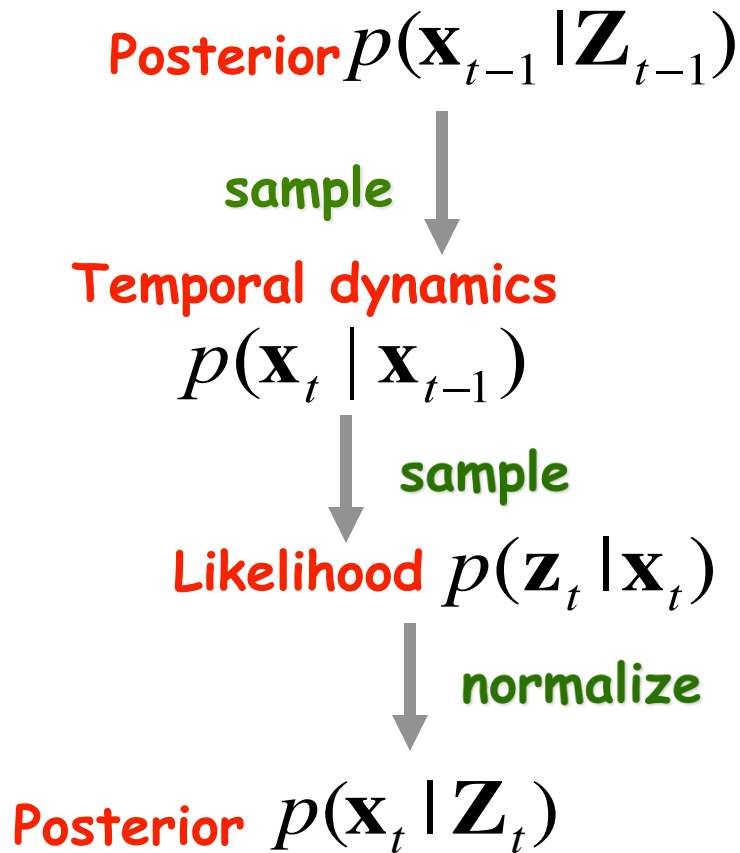
Isard & Blake '96

Particle Filter



Isard & Blake '96

Particle Filter



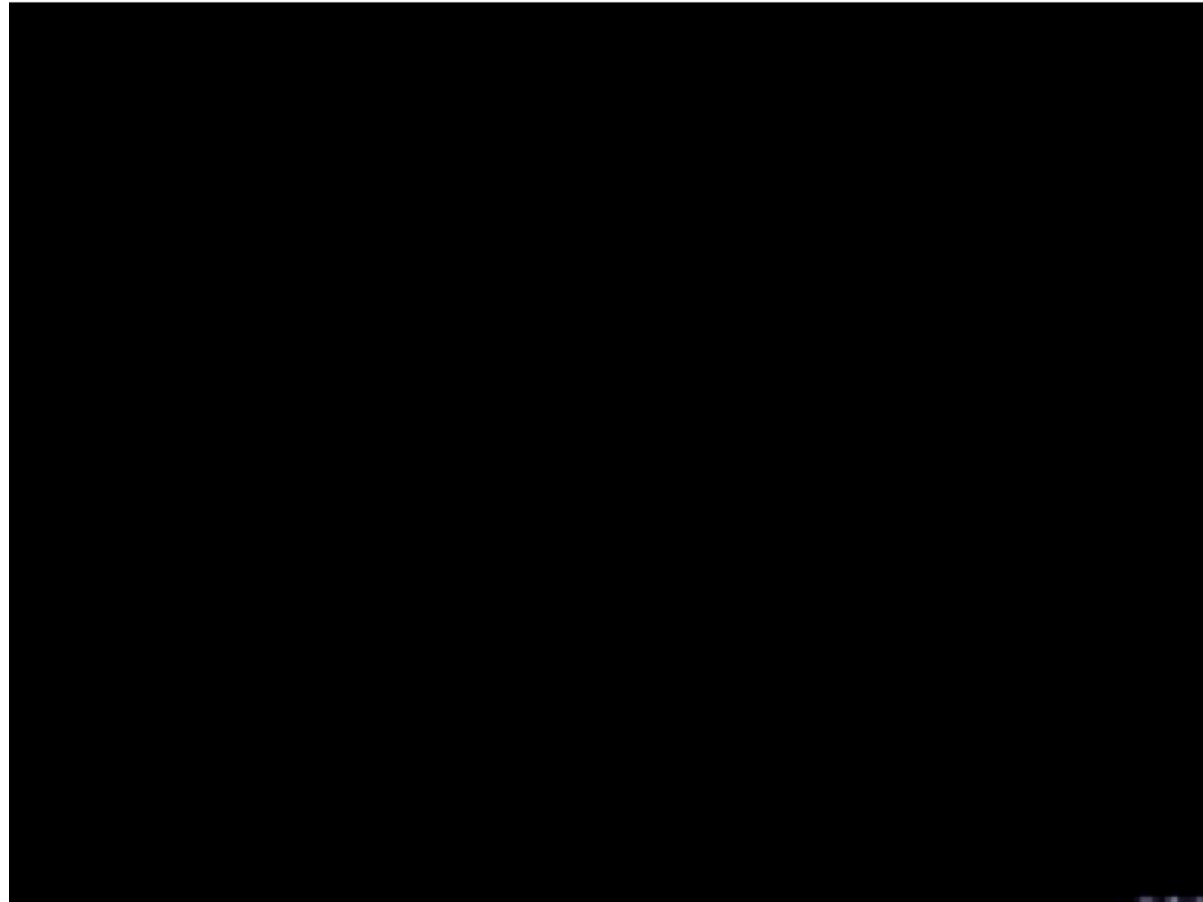
Isard & Blake '96

Pseudocode

condense1 step

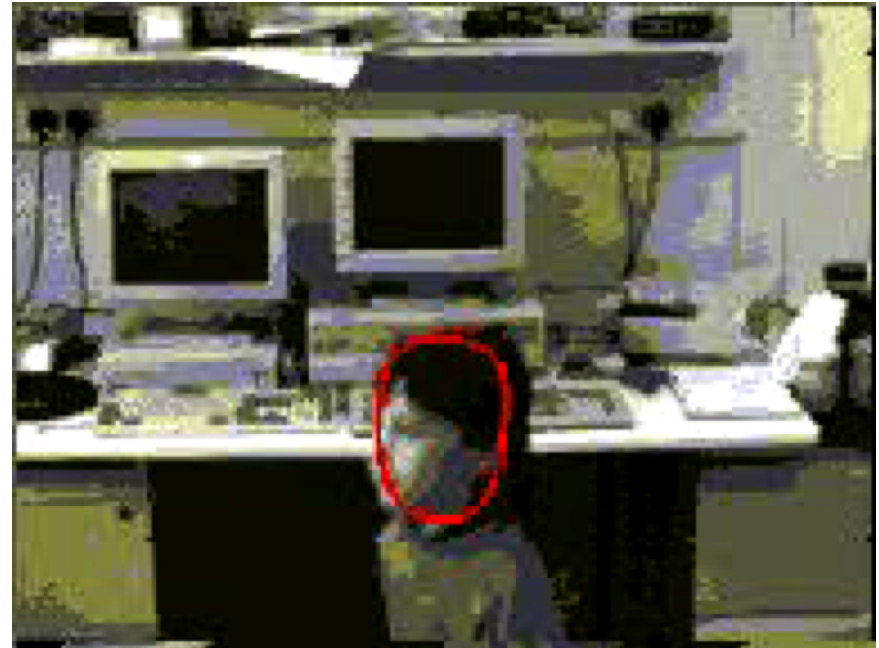
```
% generate cumulative distribution for posterior at t-1
....
% generate a vector of uniform random numbers.
% if a the number is greater than refreshRate then
    % generate a vector of uniform random numbers
    % use these to search the cumulative probability
    % find the indices of the corresponding particles
    % for each of these particles, predict the new state
    % for each of these new states compute the log likelihood
% else generate a particle at random and compute its log likelihood.
% find the maximum log likelihood and subtract it from all the other log
  likelihoods
% construct the posterior at time t by exponentiating all the log likelihoods
  and normalizing so they sum to 1.
```

Particle Filter



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Tracking in Clutter



Isard, M., Blake, and A., Condensation — Conditional density propagation for visual tracking, in *Int. J. Computer Vision*, vol. 28, no. 1, pp. 5–28, 1998