CS155/254: Probabilistic Methods in Computer Science

Chapter 13.1: Martingales

Probability and Computing

Randomization and Probabilistic Techniques in Algorithms and Data Analysis Michael Mitzenmacher and Eli Upfal SECOND EDITION

Hoeffding's Bound

Theorem

Let X_1, \ldots, X_n be independent random variables with $E[X_i] = \mu_i$ and $Pr(B_i \le X_i \le B_i + c_i) = 1$, then

$$\Pr(|\sum_{i=1}^n X_i - \sum_{i=1}^n \mu_i| \ge \epsilon) \le e^{-\frac{2\epsilon^2}{\sum_{i=1}^n c_i^2}}$$

Do we need independence?

Martingales

Definition

A sequence of random variables Z_0, Z_1, \ldots is a *martingale* with respect to the sequence X_0, X_1, \ldots if for all $n \ge 0$ the following hold:

Definition

A sequence of random variables Z_0, Z_1, \ldots is a *martingale* when it is a martingale with respect to itself, that is

• $E[|Z_n|] < \infty;$ • $E[Z_{n+1}|Z_0, Z_1, \dots, Z_n] = Z_n;$

Conditioning Defines a Probability Space

Let $(\Omega, Pr(\cdot))$ be a probability space.

Let B be an event in Ω , Pr(B) > 0.

We show that $(B, Pr(\cdot | B))$ is a probability space.

1 For any $E \subseteq B$,

$$0 \leq Pr(E \mid B) = rac{Pr(E \cap B)}{Pr(B)} \leq 1$$

2 Let E_1 and E_2 be disjoint events in B,

$$Pr(E_1 \cup E_1 \mid B) = \frac{Pr((E_1 \cup E_2) \cap B)}{Pr(B)}$$
$$= \frac{Pr(E_1 \cap B)}{Pr(B)} + \frac{Pr(E_2 \cap B)}{Pr(B)}$$
$$= Pr(E_1 \mid B) + Pr(E_2 \mid B)$$

Conditional Expectation

Definition

$$E[Y | Z = z] = \sum_{y} y Pr(Y = y | Z = z)$$
,

where the summation is over all y in the range of Y.

Note that E[Y | Z] is a random variable (a function of Z)

Lemma

For any random variables X and Y,

$$\mathsf{E}[X] = \mathsf{E}_{Y}[\mathsf{E}_{X}[X \mid Y]] = \sum_{y} \mathsf{Pr}(Y = y)\mathsf{E}[X \mid Y = y] \ ,$$

where the sum is over all values in the range of Y.

Lemma

For any random variables X and Y,

$$\mathsf{E}[X] = \mathsf{E}_{Y}[\mathsf{E}_{X}[X \mid Y]] = \sum_{y} \mathsf{Pr}(Y = y)\mathsf{E}[X \mid Y = y] ,$$

where the sum is over all values in the range of Y.

Proof.

$$\sum_{y} \Pr(Y = y) \mathbb{E}[X \mid Y = y]$$

$$= \sum_{y} \Pr(Y = y) \sum_{x} x \Pr(X = x \mid Y = y)$$

$$= \sum_{x} \sum_{y} x \Pr(X = x \mid Y = y) \Pr(Y = y)$$

$$= \sum_{x} \sum_{y} x \Pr(X = x \cap Y = y) = \sum_{x} x \Pr(X = x) = \mathbb{E}[X].$$

Example

Y - the number of students attending class, $Y \sim B(n, p)$ X - the number of questions asked in class is $X|_{Y=y} \sim B(\lfloor \sqrt{y} \rfloor, q)$. All events are independent

 $E[X \mid Y = y] = q\lfloor \sqrt{y} \rfloor - \text{ a constant}$ $E[X \mid Y] = q\lfloor \sqrt{Y} \rfloor - \text{ a random variable}$

$$\begin{split} E[X] &= E_Y[E_X[X \mid Y]] &= E_Y[q\lfloor \sqrt{Y} \rfloor] \\ &= qE[\lfloor \sqrt{Y} \rfloor] \le q\sqrt{E[Y]} = q\sqrt{np} \end{split}$$

Martingales

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Martingale Example

A series of fair games (E[gain] = 0), not necessarily independent..

Game 1: bet **\$1**.

Game i > 1: bet 2^i if won in round i - 1; bet *i* otherwise.

 X_i = amount won in *i*th game. ($X_i < 0$ if *i*th game lost).

 Z_i = total winnings at end of *i*th game.

Example

 X_i = amount won in *i*th game. ($X_i < 0$ if *i*th game lost).

 Z_i = total winnings at end of *i*th game.

 Z_1, Z_2, \ldots is martingale with respect to X_1, X_2, \ldots

 $\mathsf{E}[X_i]=0.$

 $\mathsf{E}[Z_i] = \sum_{j=1}^i \mathsf{E}[X_j] = 0 < \infty.$

 $E[Z_{i+1}|X_1, X_2, \dots, X_i] = Z_i + E[X_{i+1}] = Z_i.$

Efficient Market Hypothesis

The efficient markets hypothesis (EMH) maintains that market prices fully reflect all available information. Samuelson (1965), Fama (1963);

For simplicity assume an asset that is paying no dividend, and assume 0 interest rate (so value is not discounted in time).

Let X_t be the price of a unit asset at time t. If I know that at time t + 1 the price will be $X_{t+1} = c$, I will not sale the asset now for less than c.

If I know that at time t + 1 the price will be $X_{t+1} = c$, I will not buy the asset now for more than c.

$$X_t = E[X_{t+1} \mid X_0, \dots X_t]$$

 X_0, X_1, \ldots, X_t , is a martingale.

Gambling Strategies

I play series of fair games (win with probability 1/2).

Game 1: bet **\$1**.

Game i > 1: bet 2^i if I won in round i - 1; bet i otherwise.

 X_i = amount won in *i*th game. ($X_i < 0$ if *i*th game lost).

 Z_i = total winnings at end of *i*th game.

Assume that (before starting to play) I decide to quit after k games: what are my expected winnings?

Lemma

Let $Z_0, Z_1, Z_2, ...$ be a martingale with respect to $X_0, X_1, ...$ For any fixed n,

$\mathsf{E}_{X[0:n]}[Z_n] = \mathsf{E}_{X_0}[Z_0]$.

$$(X[0:i] = X_0,\ldots,X_i)$$

Proof.

Since Z_i is a martingale $E_{X_i}[Z_i|X_0, X_1, \dots, X_{i-1}] = Z_{i-1}$. Then

$$\mathsf{E}_{X[0:i-1]}[Z_{i-1}] = \mathsf{E}_{X[0:i-1]}[\mathsf{E}_{X_i}[Z_i|X_0, X_1, \dots, X_{i-1}]] = \mathsf{E}_{X[0:i]}[Z_i]$$

Thus,

$$E_{X[0:n]}[Z_n] = E_{X[0:n-1]}[Z_{n-1}] = \dots, = E[Z_0]$$

Gambling Strategies

I play series of fair games (win with probability 1/2).

Game 1: bet \$1.

Game i > 1: bet 2^i if I won in round i - 1; bet i otherwise.

 X_i = amount won in *i*th game. ($X_i < 0$ if *i*th game lost).

 Z_i = total winnings at end of *i*th game.

Assume that (before starting to gamble) we decide to quit after k games: what are my expected winnings?

 $\mathsf{E}[Z_k] = \mathsf{E}[Z_1] = 0.$

A Different Strategy

Same gambling game. What happens if I:

- play a random number of games?
- decide to stop only when I have won \$1000?

Stopping Time

Definition

A non-negative, integer random variable T is a stopping time for the sequence Z_0, Z_1, \ldots if the event "T = n" depends only on the value of random variables Z_0, Z_1, \ldots, Z_n .

Intuition: corresponds to a strategy for determining when to stop a sequence based only on values seen so far.

In the gambling game:

- first time I win 10 games in a row: is a stopping time;
- *the last time when I win*: is not a stopping time.

Consider again the gambling game: let T be a stopping time.

 Z_i = total winnings at end of *i*th game.

What are my winnings at the stopping time, i.e. $E[Z_T]$?

Fair game: $E[Z_k] = E[Z_0] = 0$?

"*T* =first time my total winnings are at least \$1000" is a stopping time, and $E[Z_T] > 1000...$

Martingale Stopping Theorem

Theorem

If Z_0, Z_1, \ldots is a martingale with respect to X_1, X_2, \ldots and if T is a stopping time for X_1, X_2, \ldots then (if T is finite),

$\mathsf{E}[Z_{\mathcal{T}}] = \mathsf{E}[Z_0]$

whenever one of the following holds:

- **1** there is a constant *c* such that, for all *i*, $|Z_i| \leq c$;
- **2** *T* is bounded;
- **3** $E[T] < \infty$, and there is a constant c such that $E[|Z_{i+1} Z_i||X_1, ..., X_i] < c$.

Proof of Martingale Stopping Theorem (Sketch)

Define a sequence Y_0, Y_1, \ldots such that

$$Y_i = \begin{cases} Z_i & \text{if } T > i \\ Z_T & \text{if } T \le i \end{cases}$$

Lemma

The sequence Y_0, Y_1, \ldots is a martingale with respect to Z_0, Z_1, \ldots .

Proof.

- **1** Y_n is determined by Z_0, \ldots, Z_n .
- **2** $E[|Y_n|] \le \max_{0 \le i \le n} E[|X_i|] \le \sum_{i=1}^n E[|X_i|] < \infty$
- **3** $E[Y_{n+1}|Z_0, Z_1, ..., Z_n] = Y_n + E_{Z_{n+1}}[(Y_{n+1} Y_n)1_{(T>n)}] = Y_n + E_{Z_{n+1}}[(Z_{n+1} Z_n)]Pr(T > n) = Y_n;$

Since Pr(T > n) is independent of Z_{n+1} , and $E[(Z_{n+1} - Z_n)] = 0$.

Since Y_0, Y_1, \ldots is a martingale, for any $n \ge 0$, $E[Y_n] = E[Z_0]$, and

$$\lim_{n\to\infty} E[Y_n] = E[Y_0] = E[Z_0].$$

Since T is finite, $Z_t = \lim_{n \to \infty} Z_{\min(n,T)} = \lim_{n \to \infty} Y_n$. We want to show that $E[Z_T] = \lim_{n \to \infty} E[Y_n] = E[Z_0]$.

We use a simple version of the Uniform Convergence Theorem:

Theorem

Let W_0, W_1, \ldots be a sequence of random variables such that $\lim_{n\to\infty} W_n = W$ (pointwise), and $\max_i |W_i| \le M$, where M is either a constant or a random variable with $E[|M|] < \infty$, then

 $\lim_{n\to\infty} E[W_n] = E[W].$

Proof of Martingale Stopping Theorem (Sketch)

Since *T* is finite, $\lim_{n\to\infty} Y_n = \lim_{n\to\infty} Z_{\min(n,T)} = Z_T$.

We need to show that $|Y_n| \leq M$.

- 1 there is a constant c such that, for all i, $|Z_i| \le c |Y_n| \le \max_{0 \le i \le n} |Z_i| \le c$, $c = M < \infty$.
- 2 T is bounded $|Y_n| \le \max_{0 \le i \le \max T} |Z_i| \le M < \infty$
- **3** $E[T] < \infty$, and there is a constant *c* such that $E[|Z_{i+1} Z_i||X_1, \dots, X_i] < c$

$$Y_n = Z_0 + \sum_{i=1}^{\infty} (Z_{i+1} - Z_i) \mathbf{1}_{i \le T} \le |Z_0| + \sum_{i=1}^{\infty} |Z_{i+1} - Z_i| \mathbf{1}_{i \le T} = M.$$

$$E[|M|] = E[|Z_0|] + \sum_{i=1}^{\infty} E[E[|Z_{i+1} - Z_i| | X_1, \dots, X_i] \mathbf{1}_{i \le T}]$$

$$\leq E[|Z_0|] + c \sum_{i=1}^{\infty} Pr(T \ge i)$$

$$\leq E[|Z_0|] + cE[T] < \infty$$

Martingale Stopping Theorem Applications

We play a sequence of fair game with the following stopping rules:

- **1** T is bounded, $E[Z_T] = E[Z_0]$.
- **2** T is the first time we made \$1000: E[T] is unbounded.
- **3** We double until the first win. E[T] = 2 but $E[|Z_{i+1} Z_i||X_1, ..., X_i]$ is unbounded.

Example: The Gambler's Ruin

- Consider a sequence of independent, fair 2-player gambling games.
- In each round, each player wins or loses \$1 with probability $\frac{1}{2}$.
- X_i = amount won by player 1 on *i*th round.
 - If player 1 has lost in round $i: X_i < 0$.
- Z_i = total amount won by player 1 after *i*th rounds.

• $Z_0 = 0.$

- Game ends when one player runs out of money
 - Player 1 must stop when she loses net ℓ_1 dollars $(Z_t = -\ell_1)$
 - Player 2 terminates when she loses net ℓ_2 dollars $(Z_t = \ell_2)$.
- q = probability game ends with player 1 winning ℓ_2 dollars.

Example: The Gambler's Ruin

- T = first time player 1 wins ℓ_2 dollars or loses ℓ_1 dollars.
 - T is a stopping time for X_1, X_2, \ldots
- Z_0, Z_1, \ldots is a martingale.
 - Z_i's are bounded.
- Martingale Stopping Theorem: $E[Z_T] = E[Z_0] = 0$.

$$\mathsf{E}[Z_T] = q\ell_2 - (1-q)\ell_1 = 0$$

$$q = \frac{\ell_1}{\ell_1 + \ell_2}$$

• Candidate **A** and candidate **B** run for an election.

- Candidate **A** gets *a* votes.
- Candidate **B** gets **b** votes.

• *a* > *b*.

- Votes are counted in *random order*.
 - chosen from all permutations on n = a + b votes.
- What is the probability that A is always ahead in the count?

- S_i = number of votes **A** is leading by after *i* votes counted
 - If **A** is trailing: $S_i < 0$.

•
$$S_n = a - b$$
.

• For
$$0 \le k \le n-1$$
: $X_k = \frac{S_{n-k}}{n-k}$.

• Consider X_0, X_1, \ldots, X_n .

This sequence goes backward in time!

 $E[X_k|X_0, X_1, \ldots, X_{k-1}] = ?$

 $E[X_k|X_0, X_1, \ldots, X_{k-1}] = ?$

- Conditioning on X₀, X₁,..., X_{k-1}: equivalent to conditioning on S_n, S_{n-1},..., S_{n-(k-1)},
- a_i = number of votes for **A** after first *i* votes are counted.
- (n k + 1)th vote: random vote among these first n k + 1 votes.

$$S_{n-k} = \begin{cases} S_{n-(k-1)} + 1 & \text{if } (n-k+1) \text{th vote is for } \mathbf{B} \\ S_{n-(k-1)} - 1 & \text{if } (n-k+1) \text{th vote is for } \mathbf{A} \end{cases}$$

$$S_{n-k} = \begin{cases} S_{n-(k-1)} + 1 & \text{with prob.} \quad \frac{n-k+1-a_{n-(k-1)}}{n-(k-1)} \\ S_{n-(k-1)} - 1 & \text{with prob.} \quad \frac{a_{n-(k-1)}}{n-(k-1)} \end{cases}$$

$$E[S_{n-k}|S_{n-(k-1)}] = (S_{n-(k-1)}+1)\frac{n-k+1-a_{n-(k-1)}}{(n-k+1)} + (S_{n-(k-1)}-1)\frac{a_{n-(k-1)}}{(n-k+1)} = S_{n-(k-1)}\frac{n-k}{n-(k-1)}$$

(Since
$$2a_{n-(k-1)} - n - k + 1 = a_{n-(k-1)} - b_{n-(k-1)} = S_{n-(k-1)}$$
)
(Since
 $n - k + 1 - 2a_{n-(k-1)} - b_{n-(k-1)} - a_{n-(k-1)} = -S_{n-(k-1)}$)
 $E[X_k | X_0, X_1, \dots, X_{k-1}] = E\left[\frac{S_{n-k}}{n-k} \middle| S_n, \dots, S_{n-(k-1)}\right]$
 $= \frac{S_{n-(k-1)}}{n-(k-1)}$
 $= X_{k-1}$

 \implies X_0, X_1, \ldots, X_n is a martingale.

 $T = \begin{cases} \min\{k < n-1 : X_k = 0\} & \text{if such } k \text{ exists} \\ n-1 & \text{otherwise} \end{cases}$

- T is a stopping time.
- **T** is bounded.
- Martingale Stopping Theorem:

$$\mathsf{E}[X_T] = \mathsf{E}[X_0] = \frac{\mathsf{E}[S_n]}{n} = \frac{a-b}{a+b}$$

Two cases:

- **1** A leads throughout the count.
- **2** A does not lead throughout the count.

1 A leads throughout the count. For $0 \le k \le n - 1$: $S_{n-k} > 0$, then $X_k > 0$.

T=n-1.

 $X_T = X_{n-1} = S_1.$

A gets the first vote in the count: S₁ = 1, then X_T = 1.
2 A does not lead throughout the count.
For some k: S_k = 0. Then X_k = 0.

T = k < n - 1.

 $X_T = 0.$

Putting all together:

1 A leads throughout the count: $X_T = 1$.

2 A does not lead throughout the count: $X_T = 0$

$$\mathsf{E}[X_T] = \frac{a-b}{a+b} = 1 * \mathsf{Pr}(\mathsf{Case}\ 1) + 0 * \mathsf{Pr}(\mathsf{Case}\ 2)$$

That is

$$\Pr(\mathbf{A} \text{ leads throughout the count}) = rac{a-b}{a+b}$$
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