

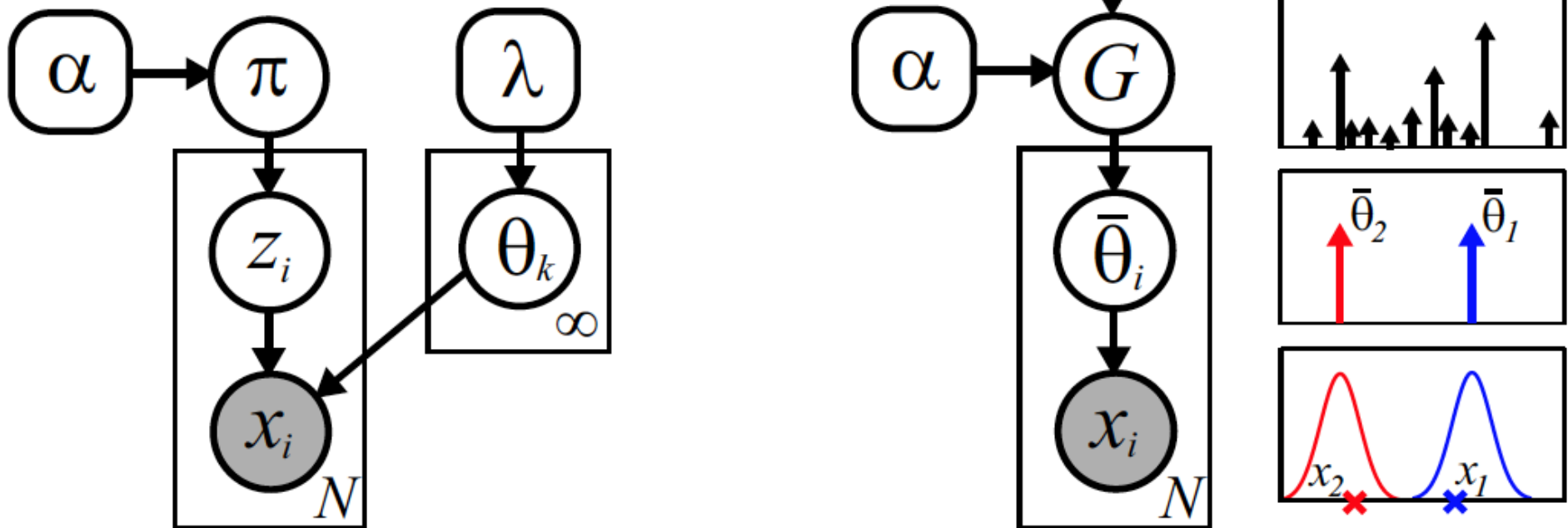
Applied Bayesian Nonparametrics

Special Topics in Machine Learning
Brown University CSCI 2950-P, Fall 2011

October 20: Binary Latent Feature Models,
Indian Buffet Process

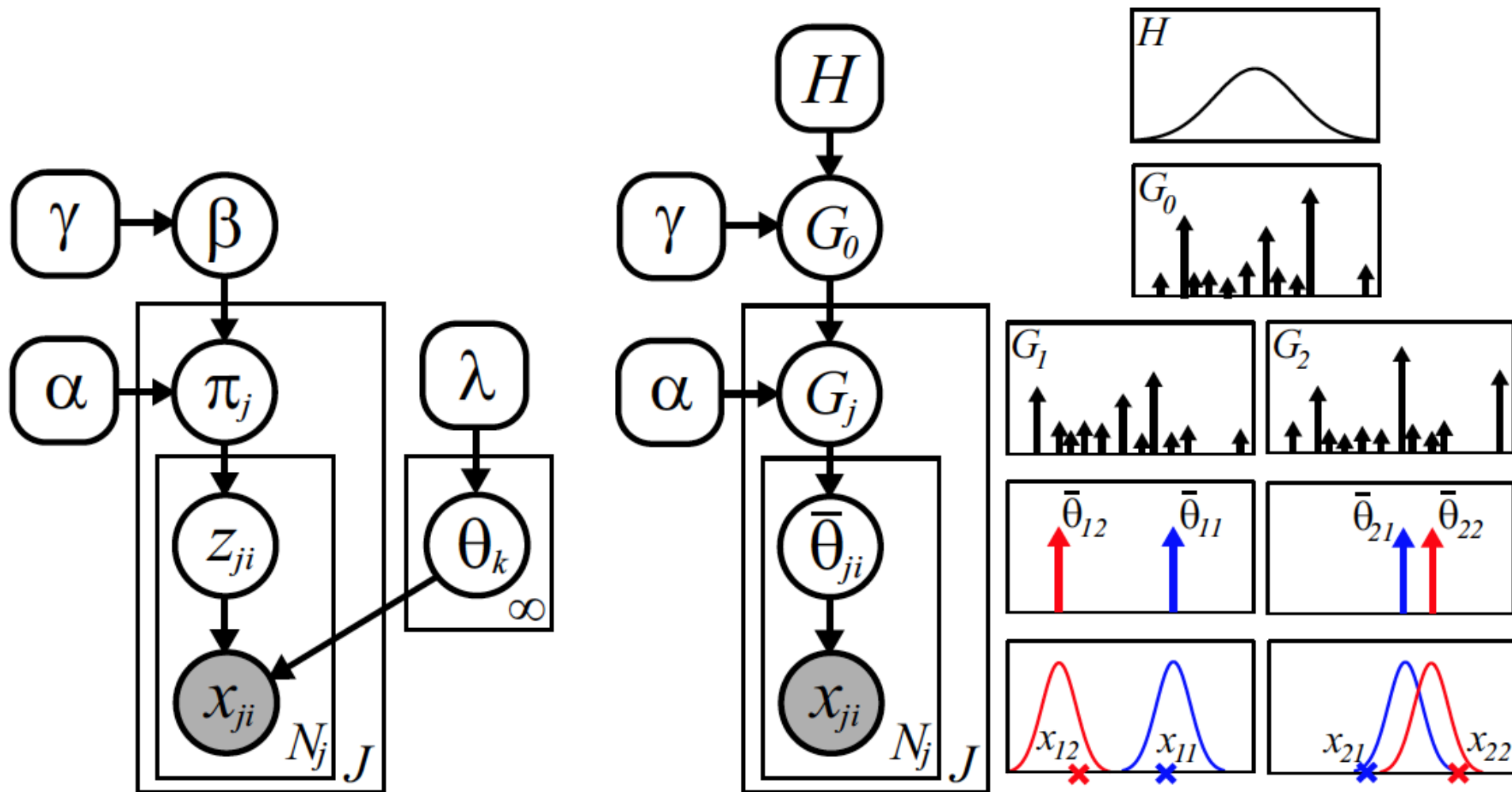
Dirichlet Process Mixtures

$$p(x | \pi, \theta_1, \theta_2, \dots) = \sum_{k=1}^{\infty} \pi_k f(x | \theta_k)$$



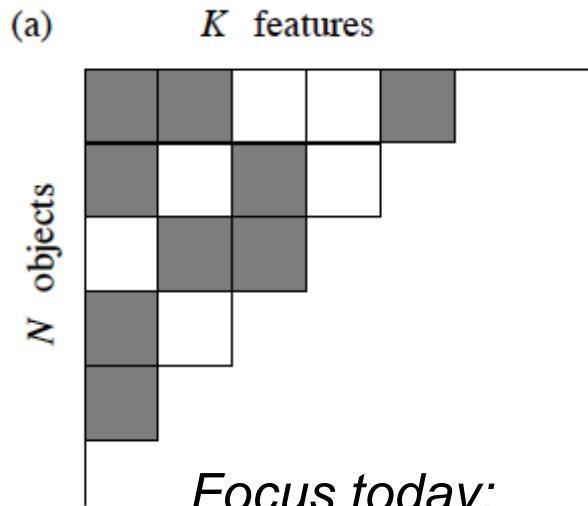
- **Mixture** model: Each observation is associated with exactly one underlying cluster or class
- Complex datasets may require a huge number of classes

Hierarchical Dirichlet Processes

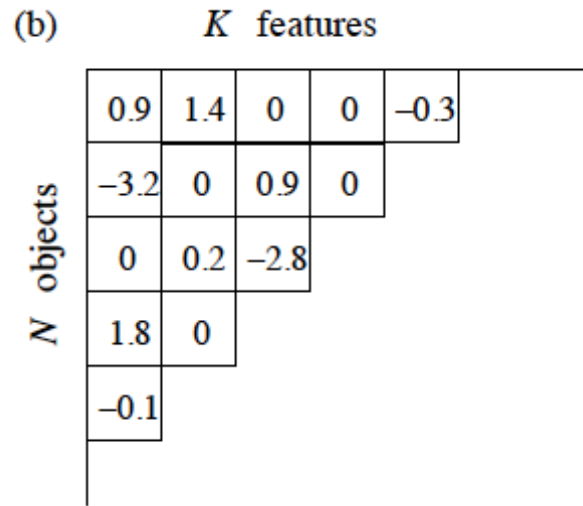


- **Admixture** model: Each group of observations is associated with a *distribution* over latent clusters/classes/topics
- Conservation: Increasing mass of one topic decreases others

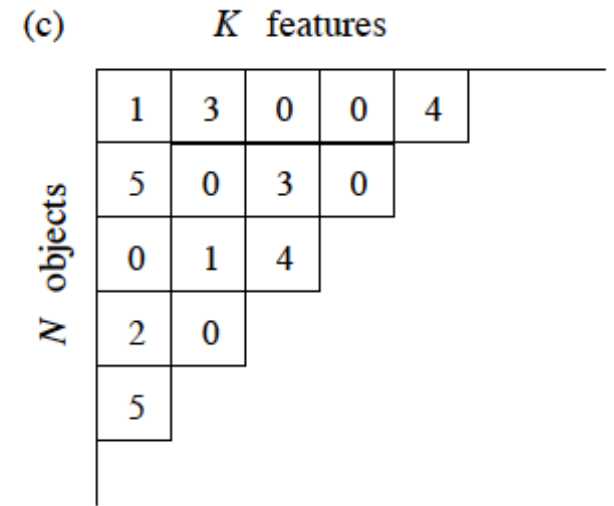
Latent Feature Models



*Focus today:
Distributions on
binary matrices
indicating feature
presence/absence*

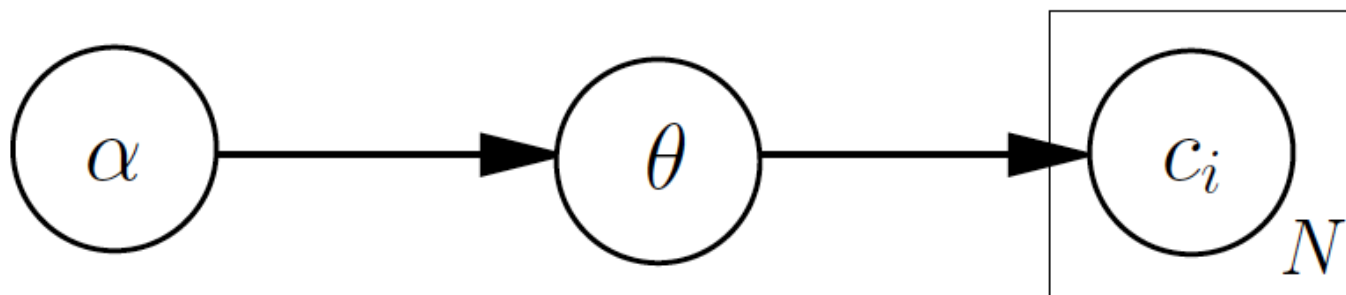


*Depending on application, features
can be associated with any
parameter value of interest*



- **Latent Feature** model: Each group of observations is associated with a *subset* of the possible latent features
- Factorial power: There are 2^K combinations of K features
- Question: What is the analog of the DP for feature modeling?

Finite Dirichlet Mixtures



$$p(\theta) = \frac{\prod_{k=1}^K \theta_k^{\alpha_k - 1}}{D(\alpha_1, \alpha_2, \dots, \alpha_K)}$$

$$P(\mathbf{c}|\theta) = \prod_{i=1}^N P(c_i|\theta) = \prod_{i=1}^N \theta_{c_i}$$

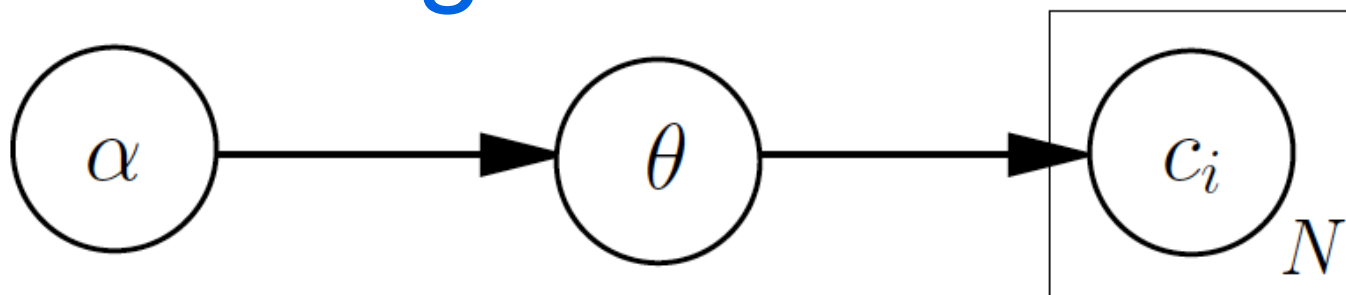
$$D(\alpha_1, \alpha_2, \dots, \alpha_K) = \int_{\Delta_K} \prod_{k=1}^K \theta_k^{\alpha_k - 1} d\theta = \frac{\prod_{k=1}^K \Gamma(\alpha_k)}{\Gamma(\sum_{k=1}^K \alpha_k)}$$

$$D\left(\frac{\alpha}{K}, \frac{\alpha}{K}, \dots, \frac{\alpha}{K}\right) = \frac{\Gamma\left(\frac{\alpha}{K}\right)^K}{\Gamma(\alpha)}$$

$$P(\mathbf{c}) = \int_{\Delta_K} \prod_{i=1}^n P(c_i|\theta) p(\theta) d\theta = \frac{\prod_{k=1}^K \Gamma(m_k + \frac{\alpha}{K})}{\Gamma\left(\frac{\alpha}{K}\right)^K} \frac{\Gamma(\alpha)}{\Gamma(N + \alpha)}$$

Marginal likelihoods generally expressed as ratios of normalizers

From Assignments to Partitions



$$\begin{aligned}
 P(\mathbf{c}) &= \int_{\Delta_K} \prod_{i=1}^n P(c_i|\theta) p(\theta) d\theta = \frac{\prod_{k=1}^K \Gamma(m_k + \frac{\alpha}{K})}{\Gamma(\frac{\alpha}{K})^K} \frac{\Gamma(\alpha)}{\Gamma(N + \alpha)} \\
 &= \left(\frac{\alpha}{K}\right)^{K_+} \left(\prod_{k=1}^{K_+} \prod_{j=1}^{m_k-1} \left(j + \frac{\alpha}{K}\right) \right) \frac{\Gamma(\alpha)}{\Gamma(N + \alpha)}
 \end{aligned}$$

K_+ is the number of classes for which $m_k > 0$.

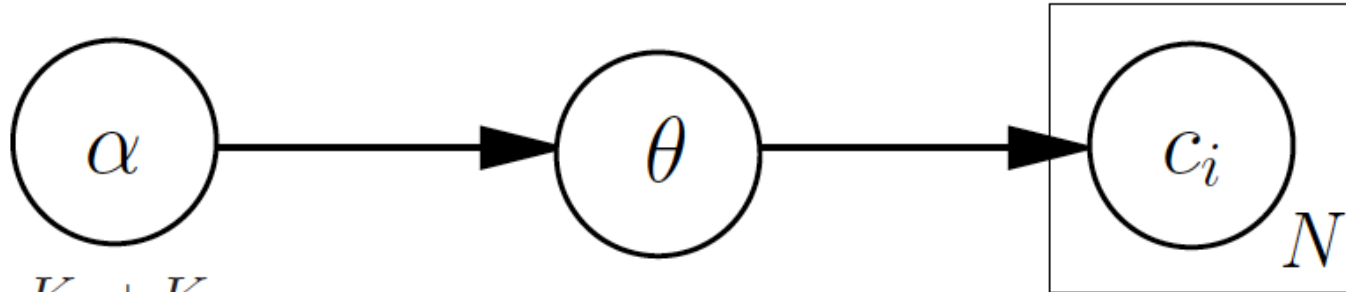
There are K^N possible values for \mathbf{c}

$P(\mathbf{c}) \rightarrow 0$ as $K \rightarrow \infty$

Instead look at label equivalence classes: $K = K_0 + K_+$

$$P([\mathbf{c}]) = \sum_{\mathbf{c} \in [\mathbf{c}]} P(\mathbf{c}) = \frac{K!}{K_0!} \left(\frac{\alpha}{K}\right)^{K_+} \left(\prod_{k=1}^{K_+} \prod_{j=1}^{m_k-1} \left(j + \frac{\alpha}{K}\right) \right) \frac{\Gamma(\alpha)}{\Gamma(N + \alpha)}$$

An Infinite Limit



$$K = K_0 + K_+$$

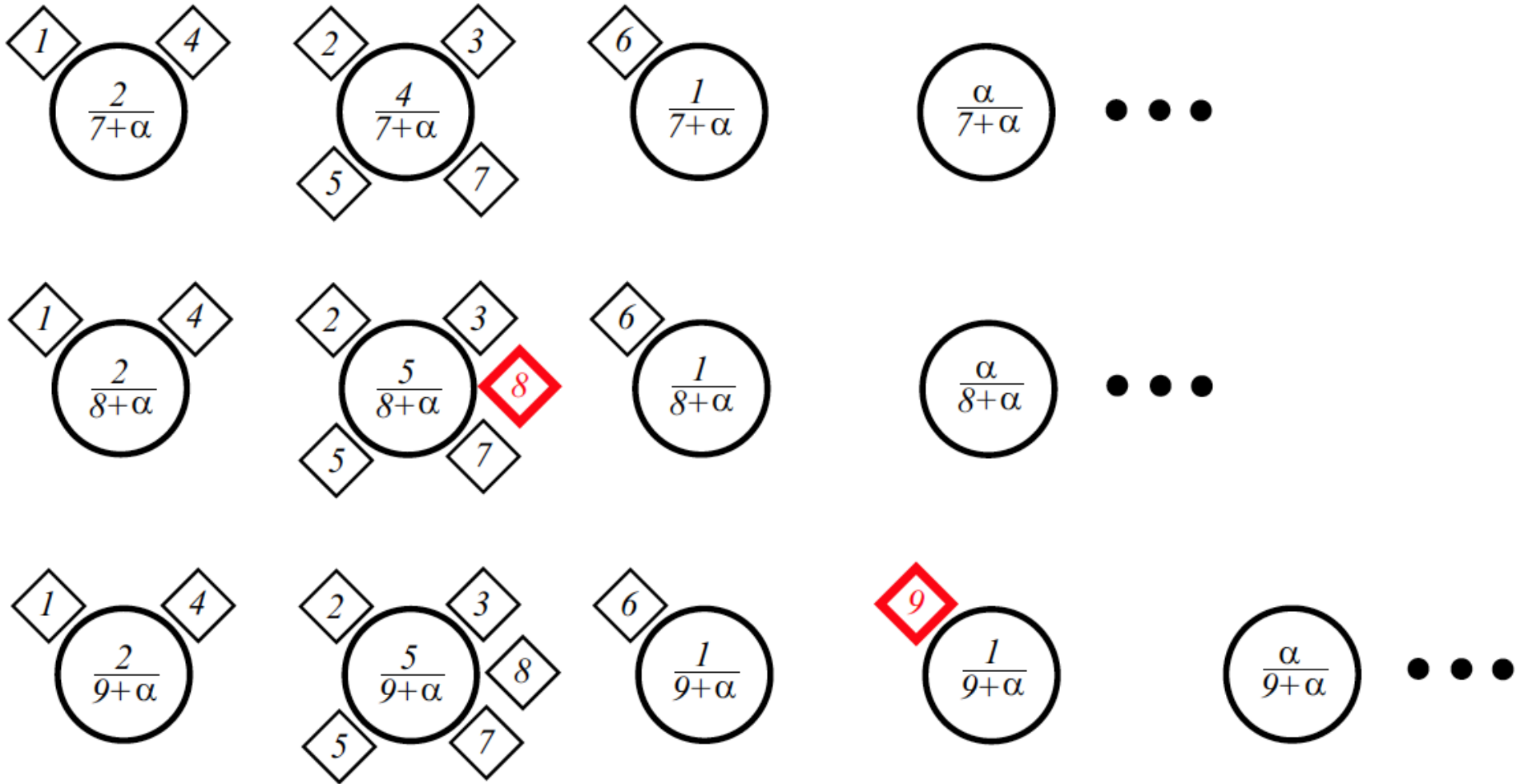
$$P([\mathbf{c}]) = \sum_{\mathbf{c} \in [\mathbf{c}]} P(\mathbf{c}) = \frac{K!}{K_0!} \left(\frac{\alpha}{K}\right)^{K_+} \left(\prod_{k=1}^{K_+} \prod_{j=1}^{m_k-1} \left(j + \frac{\alpha}{K}\right) \right) \frac{\Gamma(\alpha)}{\Gamma(N + \alpha)}$$

$$\lim_{K \rightarrow \infty} \alpha^{K_+} \cdot \frac{K!}{K_0! K^{K_+}} \cdot \left(\prod_{k=1}^{K_+} \prod_{j=1}^{m_k-1} \left(j + \frac{\alpha}{K}\right) \right) \cdot \frac{\Gamma(\alpha)}{\Gamma(N + \alpha)}$$

$$= \alpha^{K_+} \cdot 1 \cdot \left(\prod_{k=1}^{K_+} (m_k - 1)! \right) \cdot \frac{\Gamma(\alpha)}{\Gamma(N + \alpha)}$$

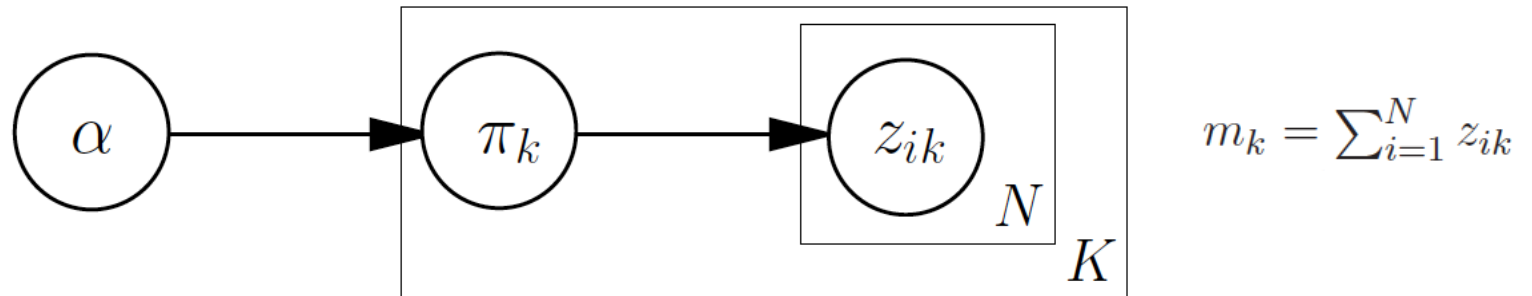
$$\begin{aligned} \frac{K!}{K_0! K^{K_+}} &= \frac{\prod_{k=1}^{K_+} (K - k + 1)}{K^{K_+}} \\ &= \frac{K^{K_+} - \frac{(K_+-1)K_+}{2} K^{K_+-1} + \dots + (-1)^{K_+-1} (K_+ - 1)! K}{K^{K_+}} \\ &= 1 - \frac{(K_+ - 1)K_+}{2K} + \dots + \frac{(-1)^{K_+-1} (K_+ - 1)!}{K^{K_+-1}}. \end{aligned}$$

Chinese Restaurant Process



$$p(z_{N+1} = z \mid z_1, \dots, z_N, \alpha) = \frac{1}{\alpha + N} \left(\sum_{k=1}^K N_k \delta(z, k) + \alpha \delta(z, \bar{k}) \right)$$

Finite Beta-Bernoulli Features



$$p(\pi_k) = \frac{\pi_k^{r-1} (1 - \pi_k)^{s-1}}{B(r, s)} \quad P(\mathbf{Z}|\pi) = \prod_{k=1}^K \prod_{i=1}^N P(z_{ik}|\pi_k) = \prod_{k=1}^K \pi_k^{m_k} (1 - \pi_k)^{N-m_k}$$

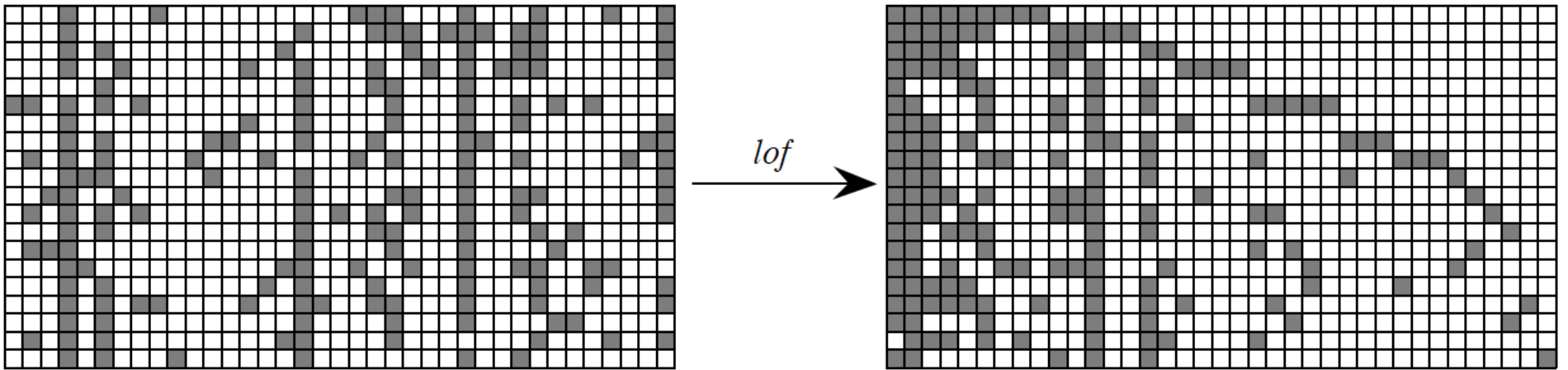
$$B(r, s) = \int_0^1 \pi_k^{r-1} (1 - \pi_k)^{s-1} d\pi_k = \frac{\Gamma(r)\Gamma(s)}{\Gamma(r+s)}$$

$$B\left(\frac{\alpha}{K}, 1\right) = \frac{\Gamma\left(\frac{\alpha}{K}\right)}{\Gamma\left(1 + \frac{\alpha}{K}\right)} = \frac{K}{\alpha}$$

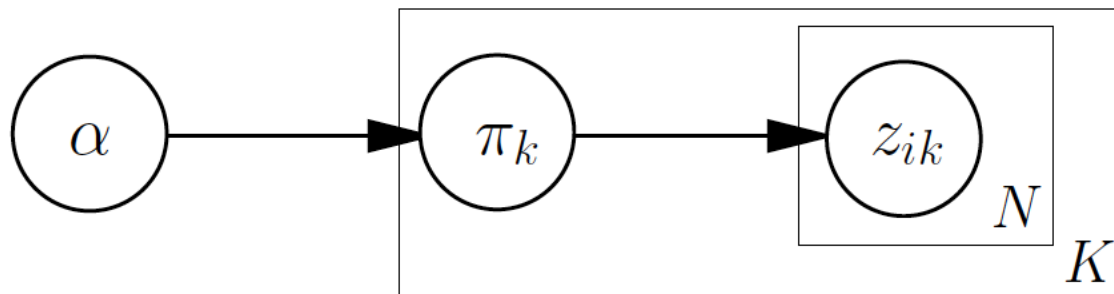
$$P(\mathbf{Z}) = \prod_{k=1}^K \int \left(\prod_{i=1}^N P(z_{ik}|\pi_k) \right) p(\pi_k) d\pi_k = \prod_{k=1}^K \frac{\frac{\alpha}{K} \Gamma(m_k + \frac{\alpha}{K}) \Gamma(N - m_k + 1)}{\Gamma(N + 1 + \frac{\alpha}{K})}$$

Marginal likelihoods generally expressed as ratios of normalizers

Left-Ordered Forms (LOFs)



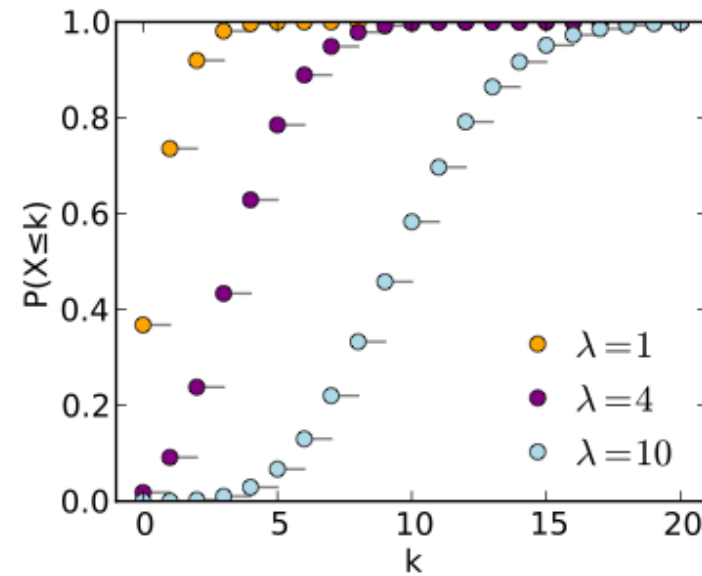
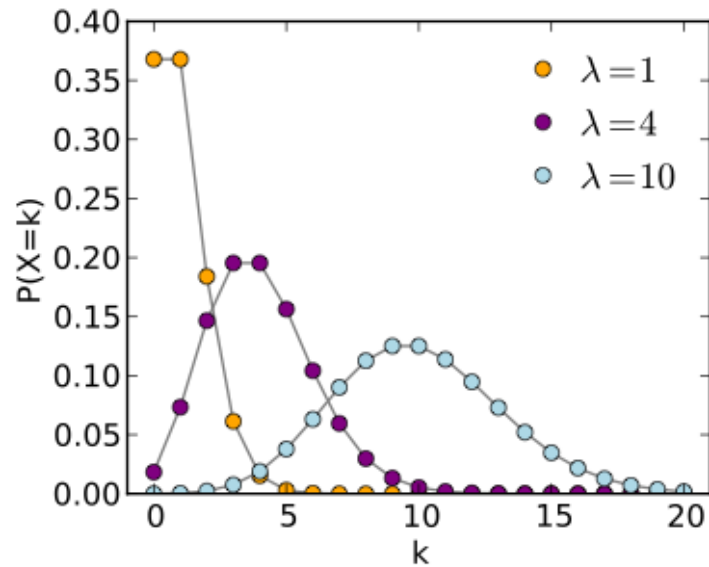
Assignments to LOFs and a Limit



Consider histories h : All possible usage patterns for one feature across the N objects.

$$\begin{aligned}
 P([\mathbf{Z}]) &= \sum_{\mathbf{Z} \in [\mathbf{Z}]} P(\mathbf{Z}) = \frac{K!}{\prod_{h=0}^{2^N-1} K_h!} \prod_{k=1}^K \frac{\frac{\alpha}{K} \Gamma(m_k + \frac{\alpha}{K}) \Gamma(N - m_k + 1)}{\Gamma(N + 1 + \frac{\alpha}{K})} \\
 \lim_{K \rightarrow \infty} &\frac{\alpha^{K_+}}{\prod_{h=1}^{2^N-1} K_h!} \cdot \frac{K!}{K_0! K^{K_+}} \cdot \left(\frac{N!}{\prod_{j=1}^N (j + \frac{\alpha}{K})} \right)^K \cdot \prod_{k=1}^{K_+} \frac{(N - m_k)! \prod_{j=1}^{m_k-1} (j + \frac{\alpha}{K})}{N!} \\
 &= \frac{\alpha^{K_+}}{\prod_{h=1}^{2^N-1} K_h!} \cdot 1 \cdot \exp\{-\alpha H_N\} \cdot \prod_{k=1}^{K_+} \frac{(N - m_k)! (m_k - 1)!}{N!}, \\
 H_N &= \sum_{j=1}^N \frac{1}{j}.
 \end{aligned}$$

Poisson Distribution



$$p(x | \lambda) = \frac{\lambda^x e^{-\lambda}}{x!}$$

$$\lambda = \mathbb{E}[x] = \mathbb{V}[x]$$

If $X_i \sim \text{Pois}(\lambda_i)$ follow a Poisson distribution with parameter λ_i and X_i are independent, then

$$Y = \sum_{i=1}^N X_i \sim \text{Pois} \left(\sum_{i=1}^N \lambda_i \right)$$

Why Poisson? The Law of Rare Events

From Wikipedia:

We will prove that, for fixed λ , if

$$X_n \sim B(n, \lambda/n); \quad Y \sim \text{Pois}(\lambda).$$

then for each fixed k

$$\lim_{n \rightarrow \infty} P(X_n = k) = P(Y = k).$$

To see the connection with the above discussion, for any Binomial random variable with large n and small p set $\lambda = np$. Note that the expectation $E(X_n) = \lambda$ is fixed with respect to n .

First, recall from [calculus](#)

$$\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n = e^{-\lambda},$$

then since $p = \lambda / n$ in this case, we have

$$\begin{aligned} \lim_{n \rightarrow \infty} P(X_n = k) &= \lim_{n \rightarrow \infty} \binom{n}{k} p^k (1-p)^{n-k} \\ &= \lim_{n \rightarrow \infty} \frac{n!}{(n-k)!k!} \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{n-k} \\ &= \lim_{n \rightarrow \infty} \underbrace{\left[\frac{n!}{n^k (n-k)!}\right]}_{A_n} \left(\frac{\lambda^k}{k!}\right) \underbrace{\left(1 - \frac{\lambda}{n}\right)^n}_{\rightarrow \exp(-\lambda)} \underbrace{\left(1 - \frac{\lambda}{n}\right)^{-k}}_{\rightarrow 1} \\ &= \left[\lim_{n \rightarrow \infty} A_n\right] \left(\frac{\lambda^k}{k!}\right) \exp(-\lambda) \end{aligned}$$

Like finite feature model, but there probabilities are random...